HW2. Start early. Otherwise you can't finish

FOR LOOP

Parallel (broadcasting) numpy



Emojis are from Pinterest.

Machine Learning CS 4641-7641

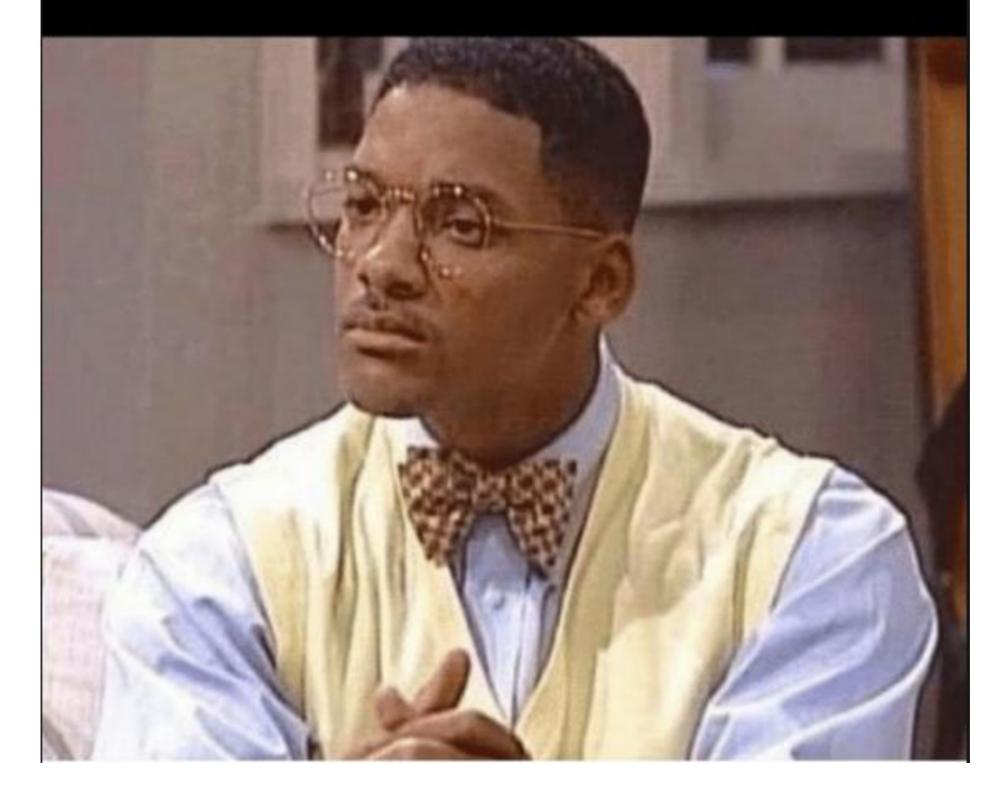


Gaussian Mixture Model

Mahdi Roozbahani Georgia Tech

Some of the slides are inspired based on slides from Jiawei Han Chao Zhang, and Barnabás Póczos.

When you're actually paying attention in class and you still have no idea what's going on

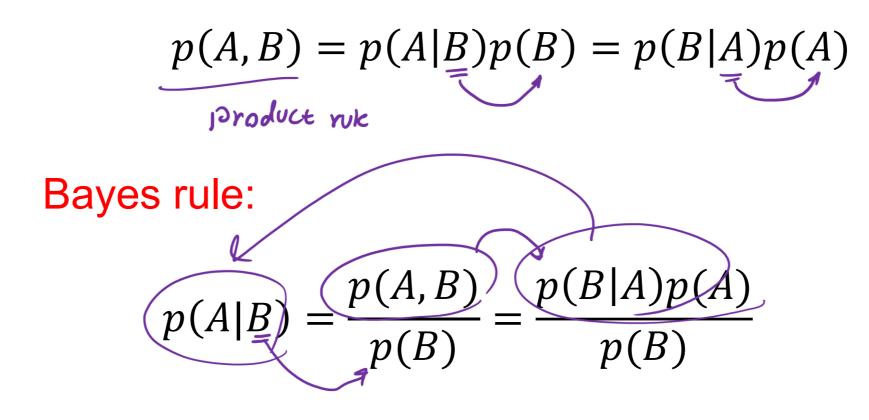


Outline

- Overview -
- Gaussian Mixture Model
- The Expectation-Maximization Algorithm

Recap

Conditional probabilities:



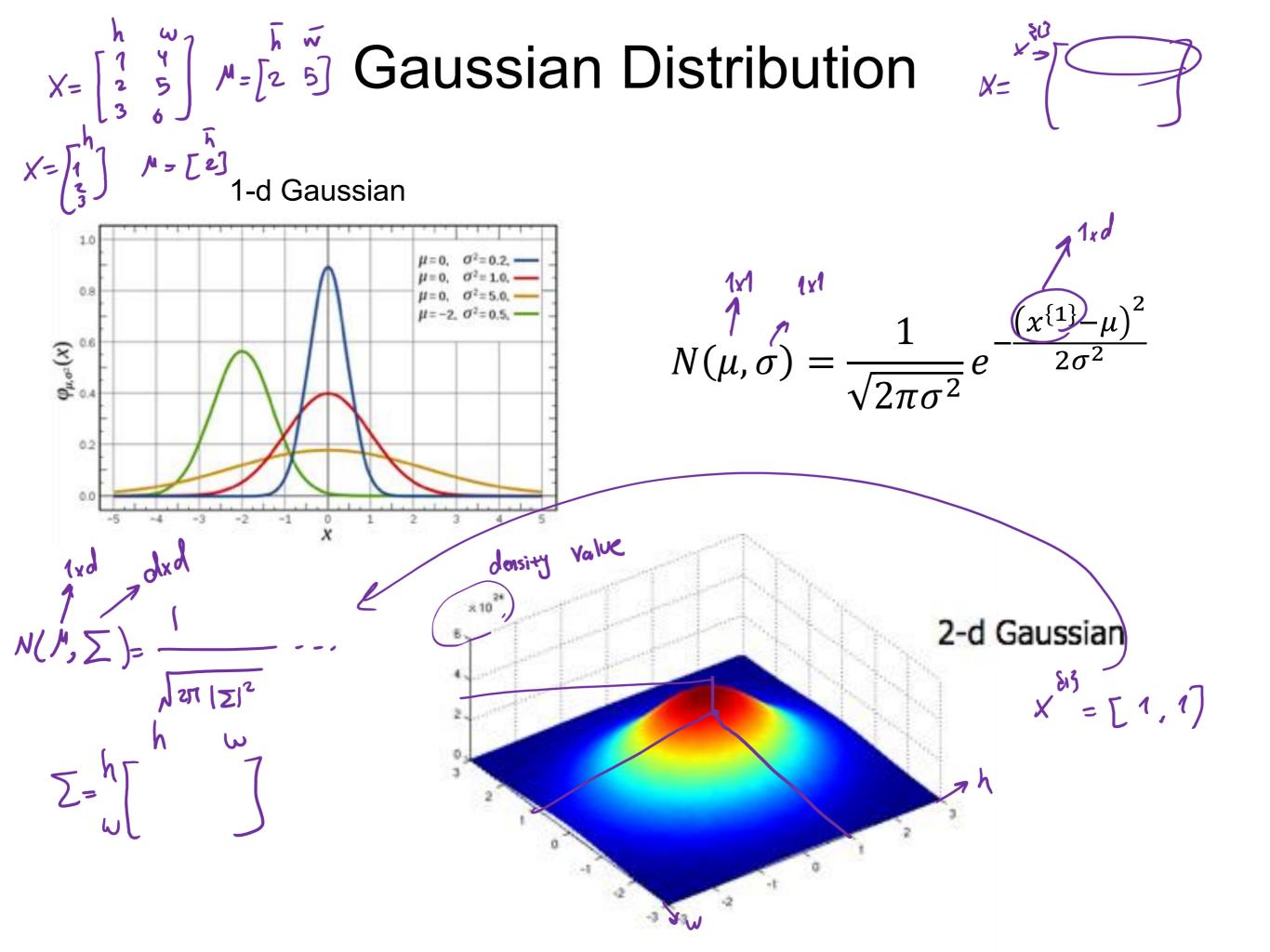
 $p(A = 1) = \sum_{i=1}^{K} p(A = 1, B_i) = \sum_{i=1}^{K} p(A|B_i) p(B_i)$

Tod	
-100	

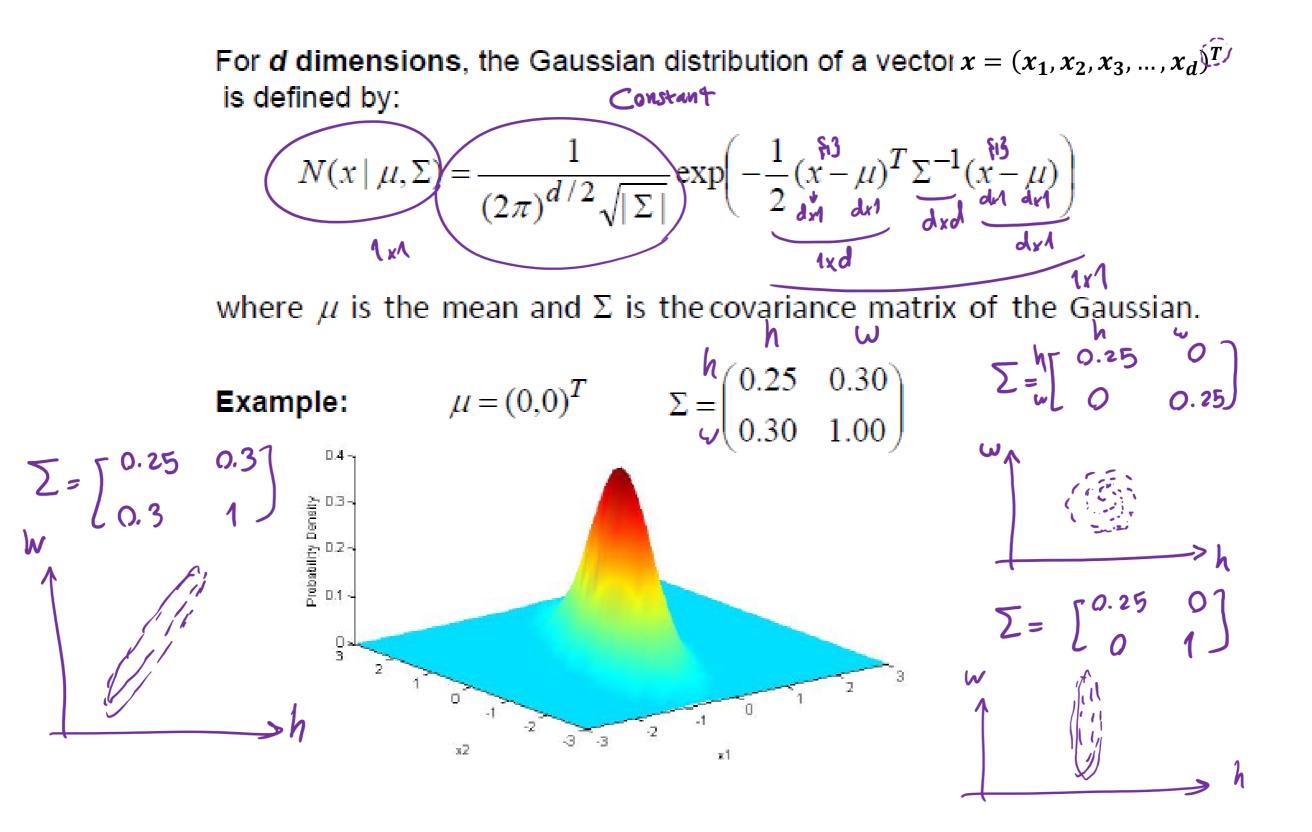
	Tomorrow=Rainy	Tomorrow=Cold	P(Today)
Today=Rainy	4/9	2/9	[4/9 + 2/9] = 2/3
Today=Cold	2/9	1/9	[2/9 + 1/9] = 1/3
P(Tomorrow)	[4/9 + 2/9] = 2/3	[2/9 + 1/9] = 1/3	

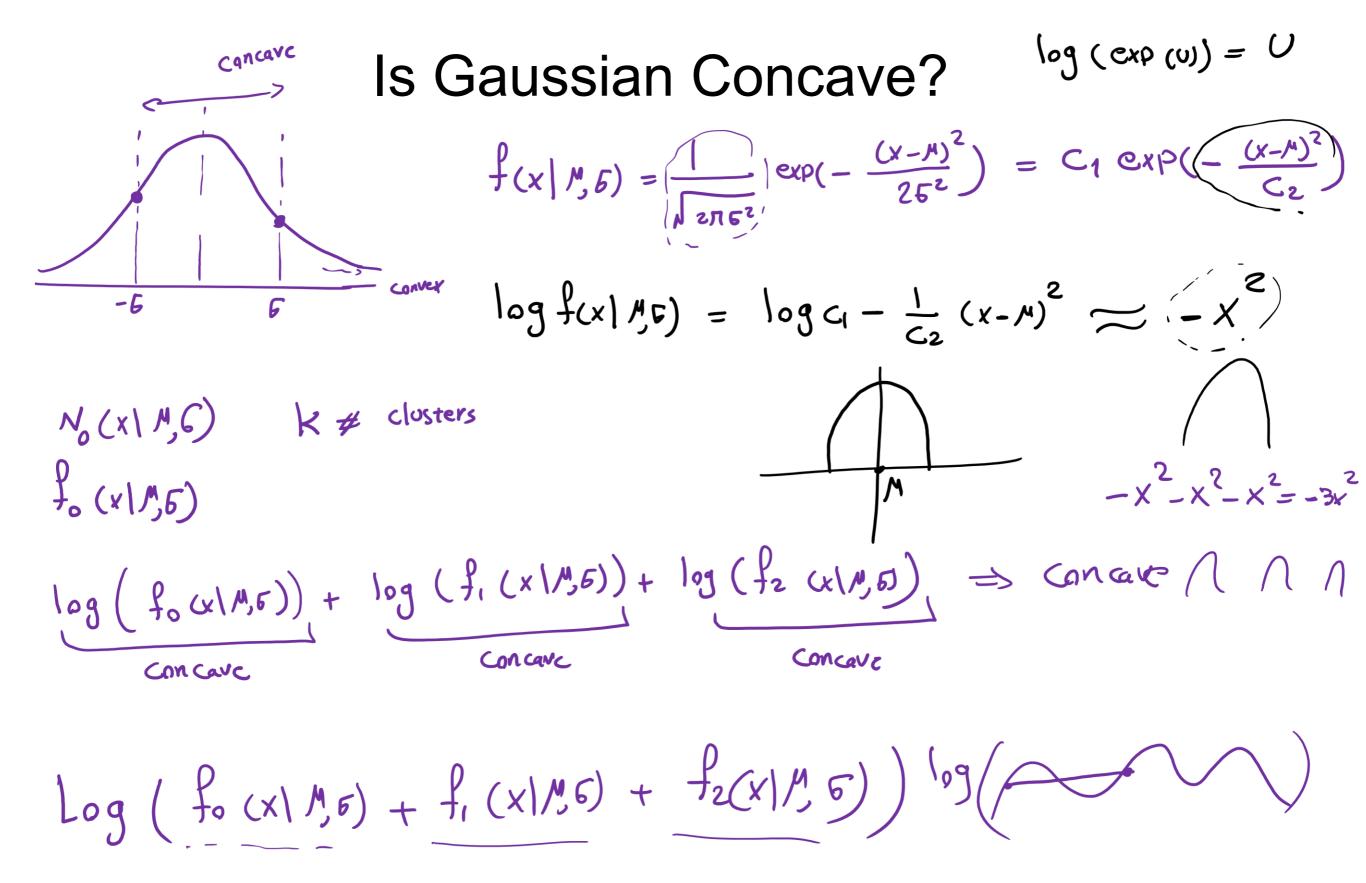
$$P(\text{Tomorrow} = \text{Rainy}) = \int_{T_0 d} P(T_{om} = \text{rainy}, T_{od}) = T_{od}$$

$$= P(\text{Tom} = \text{rainy}, \text{Tod} = \text{rainy}) + P(\text{Tom} = \text{rainy}, \text{Tod} = \text{cold})$$
$$= \frac{4}{9} + \frac{2}{9} = \frac{6}{9}$$



What is a Gaussian?



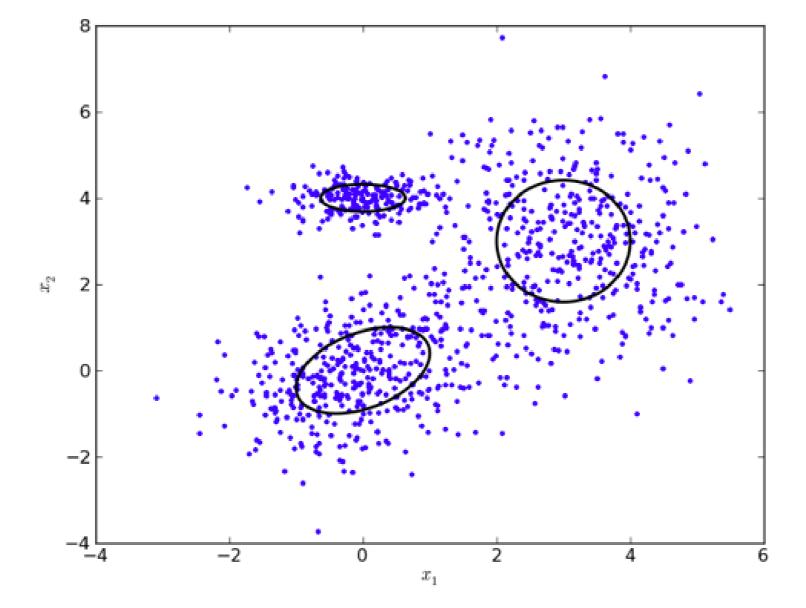


Outline

- Overview
- Gaussian Mixture Model
- The Expectation-Maximization Algorithm

Hard Clustering Can Be Difficult

• Hard Clustering: K-Means, Hierarchical Clustering, DBSCAN



Towards Soft Clustering

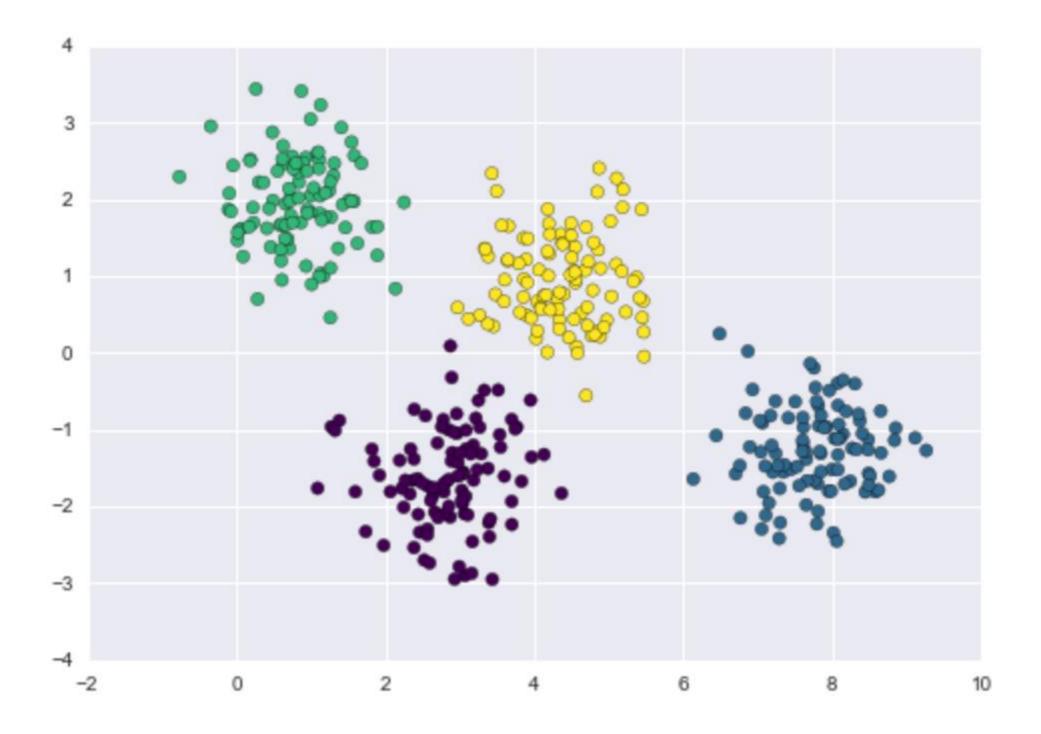
K-means

-hard assignment: each object belongs to only one cluster

$$\theta_i \in \{\theta_1, \ldots, \theta_K\}$$

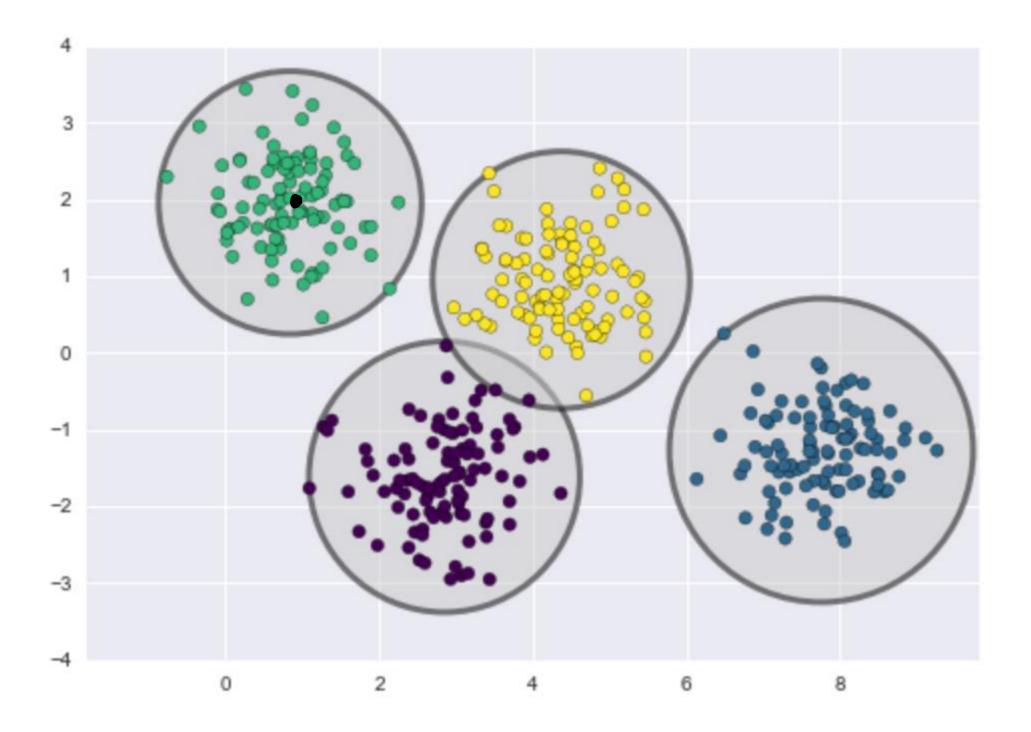
Mixture modeling

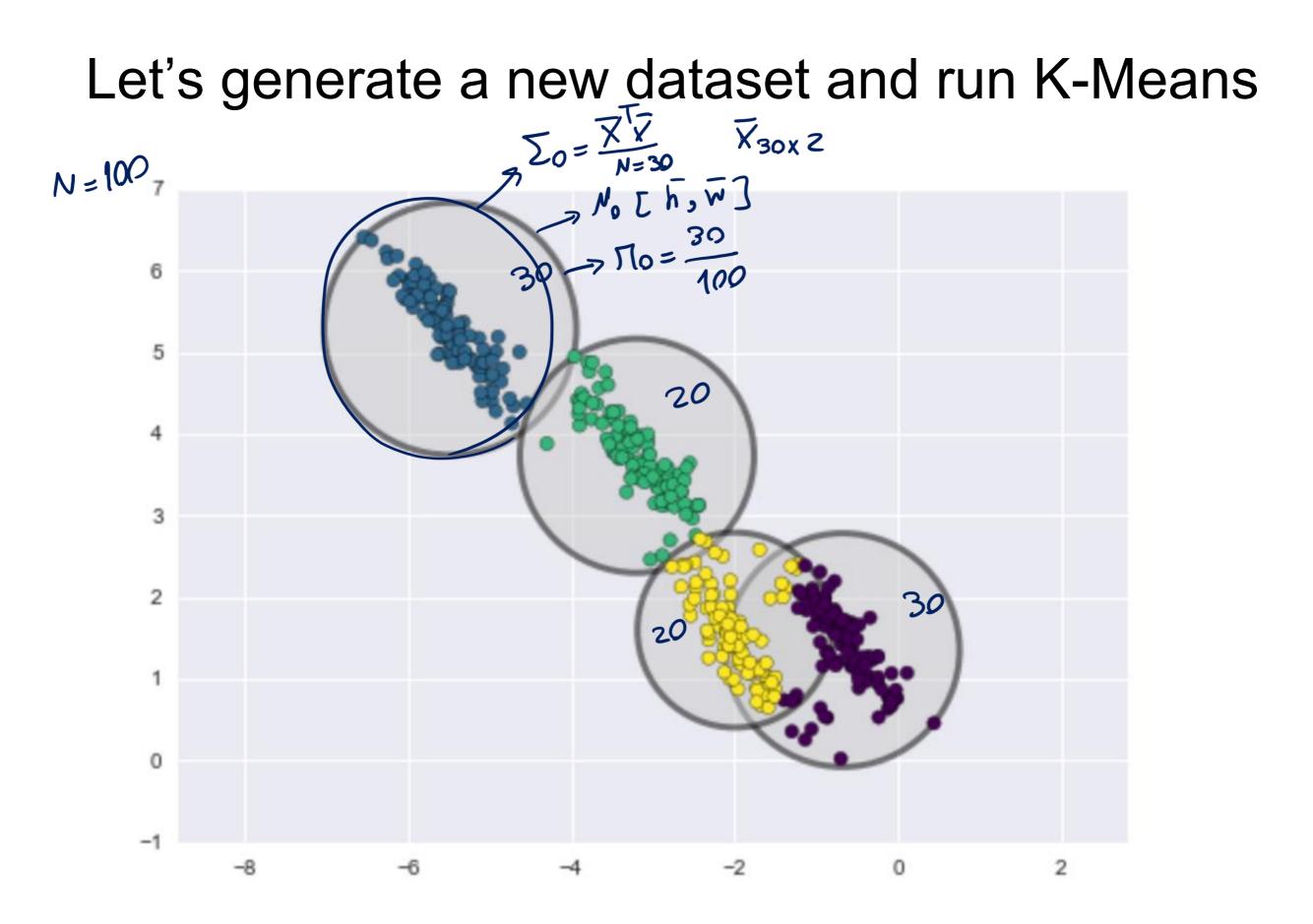
-soft assignment: probability that an object belongs to a cluster $\begin{bmatrix} 0.33 & 0.33 & 0.33 \end{bmatrix} \begin{bmatrix} 13 & 12 & 6 \\ 0.7 & (0.7) \end{bmatrix} = 1$

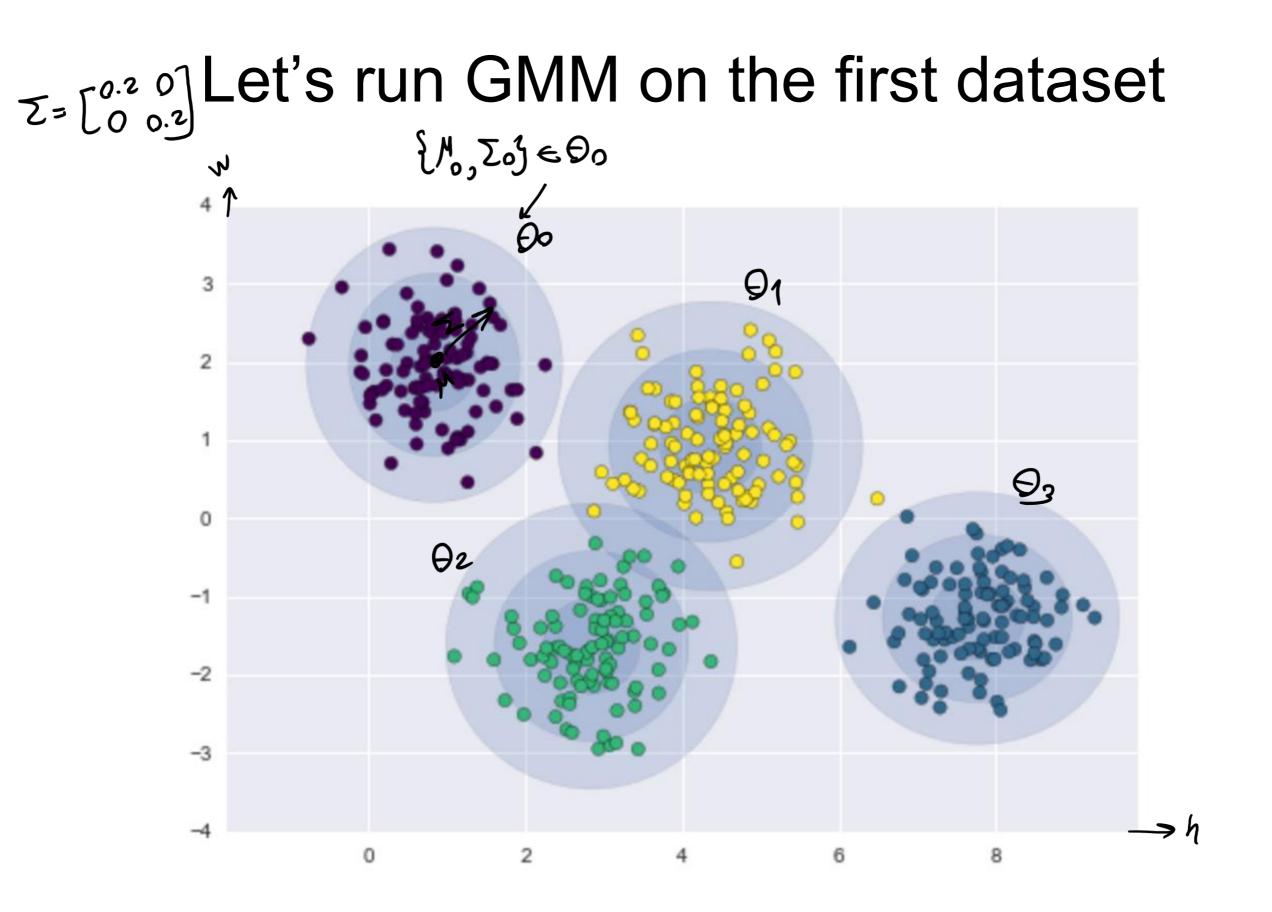


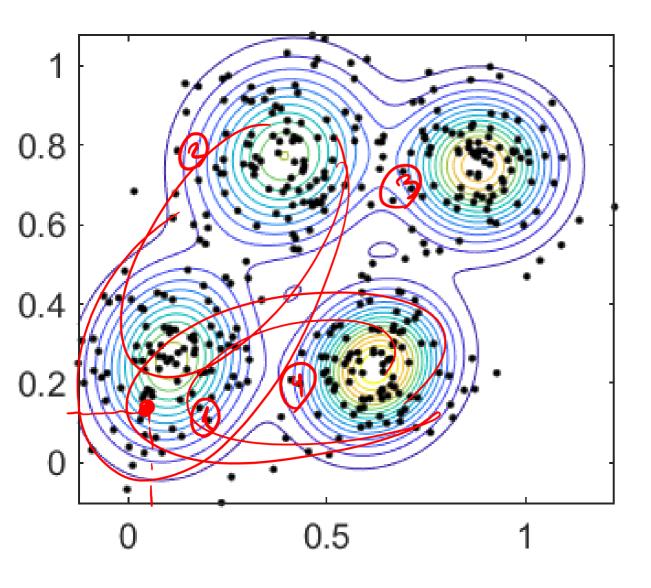
This is an excerpt from the <u>Python Data Science Handbook</u> by Jake VanderPlas

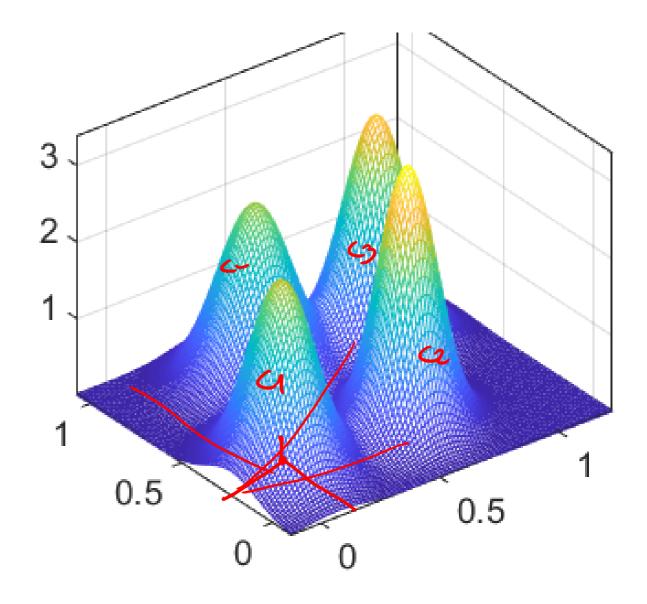
Let's run K-Means on the dataset

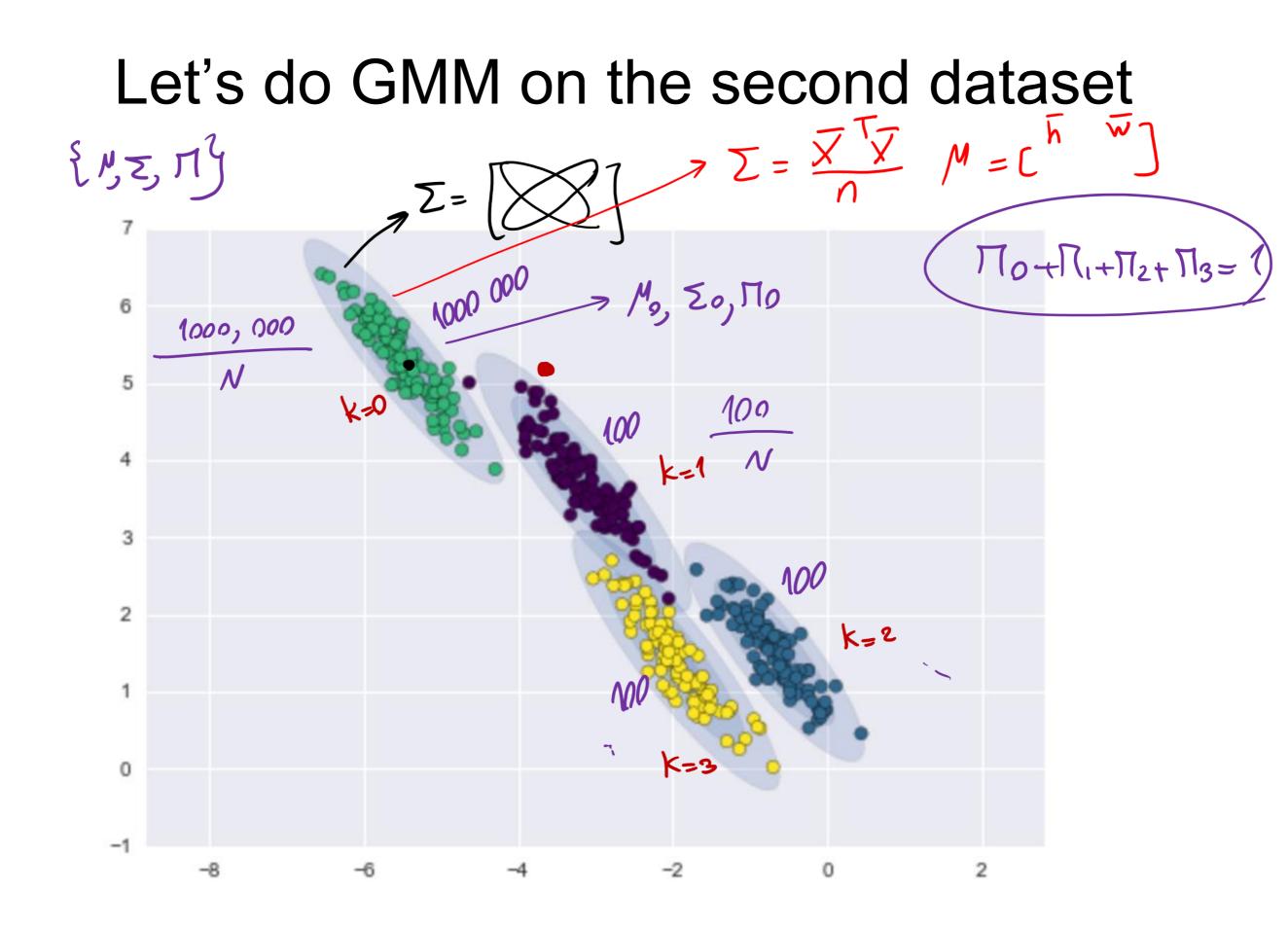


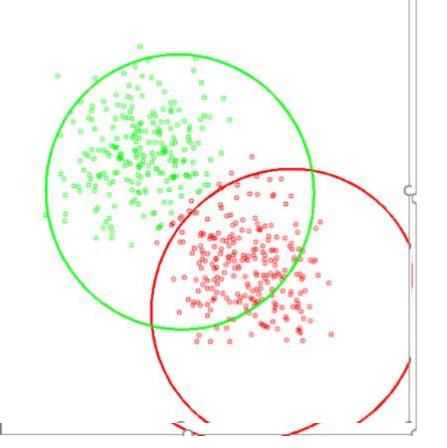


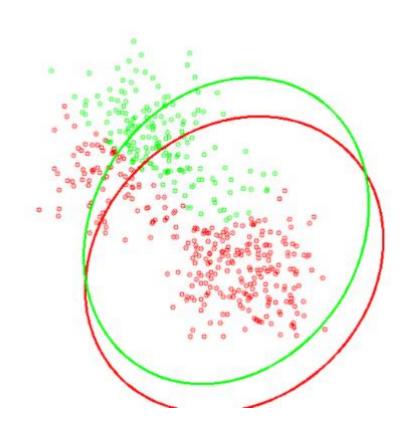


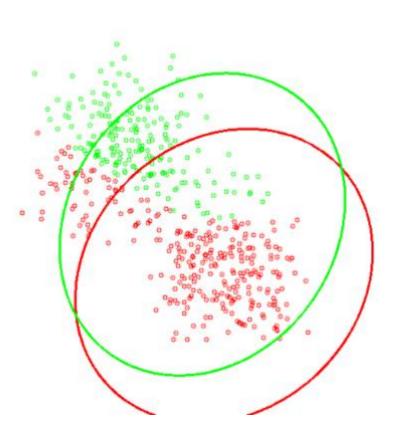


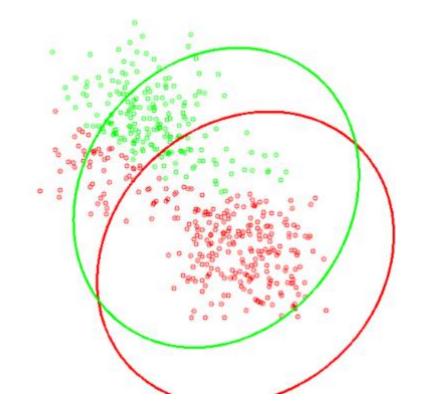


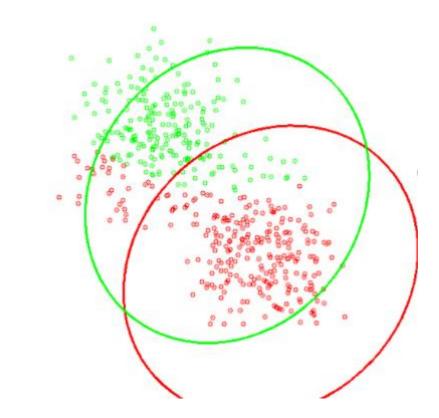


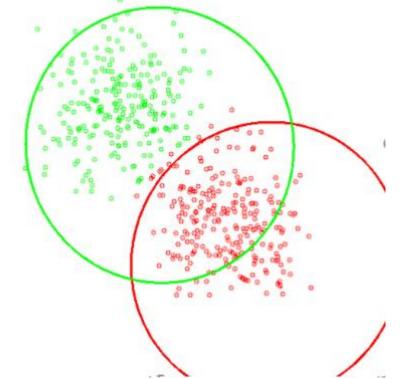




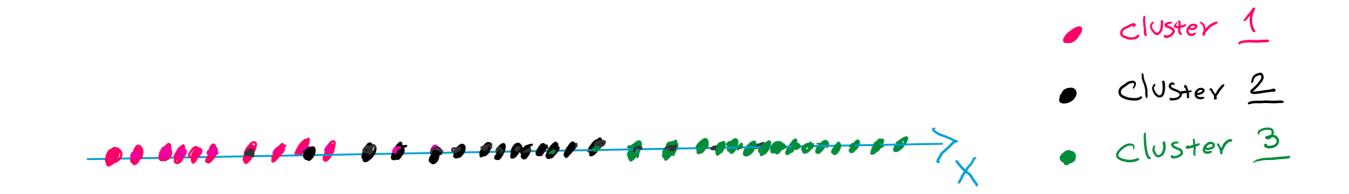






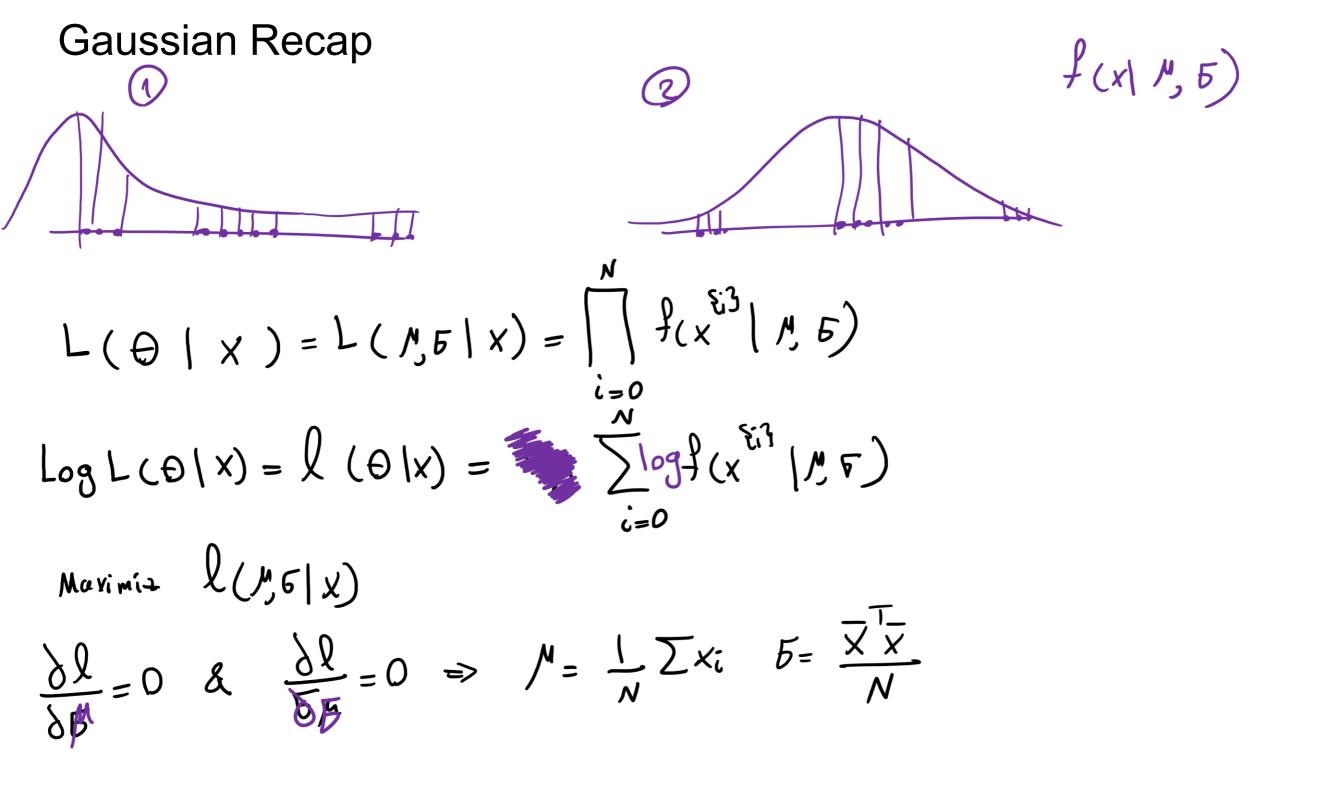


ODDIGHT FORDER BOTTOFFE A FORTOFOTIES



Initial Step Ð 2 9 ODDARD I Phil D & BO PINIP P

Final Step



Mixture perspective – soft assignment k=3 $\Sigma \pi_{i=1}$

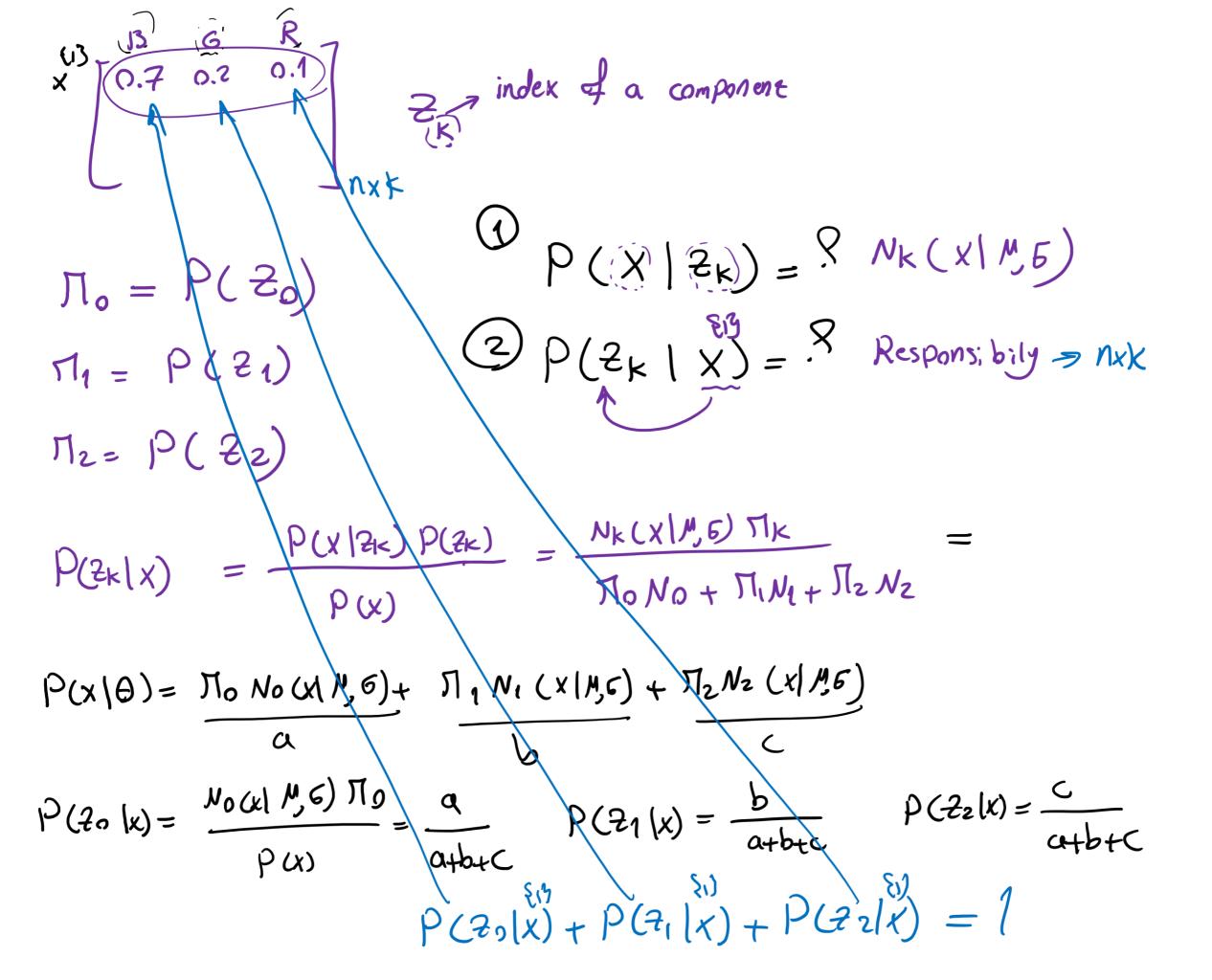
Let's create a **SINGLE** pdf that combines all three Gaussians!!!!

$$L(\Theta | x) = L(M, \varepsilon, \pi | x) = \prod_{i=1}^{N} f(x^{i3} | M, \varepsilon, \pi)$$

$$\log L(\Theta | x) = \hat{I}(\Theta | x) = \sum_{i=1}^{N} \log f(x^{i3} | M, \varepsilon, \pi)$$

$$\hat{I}(\Theta | x) = \sum_{i=1}^{N} \log \left[\Pi_0 N_0(x^{i3} | M, \varepsilon) + \Pi_1 N_1(x^{i3} | M, \varepsilon) + \Pi_2 N_2(x^{i23} | M, \varepsilon) \right]$$

Max Q(Olx)

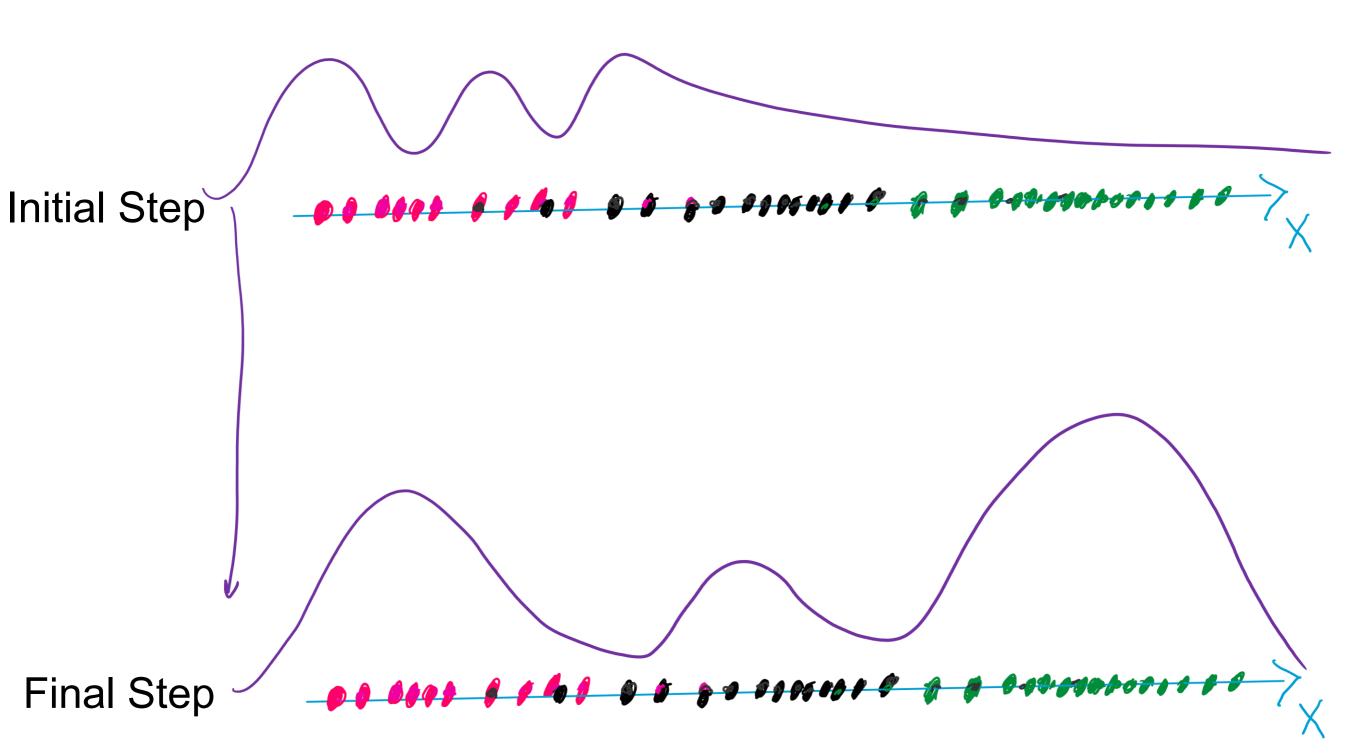


$$\begin{split} \begin{pmatrix} P(x) \\ P(x) \end{pmatrix} = P(x|\theta) = P(x|H, \Sigma, \Pi) &= \prod_{0}^{1} N_{0} \underbrace{(x|H, \Sigma_{0})}_{2} + \cdots + \prod_{k}^{n} N_{k} \underbrace{(x|H, \Sigma_{k})}_{2} \\ P(a,b) = P(a|b) P(b) \\ P(a) = \sum_{k}^{n} P(a,b_{k}) \\ P(a) = \sum_{k}^{n} P(a,b_{k}) \\ \Pi_{0} = P(Z_{0}) = P(Z_{0}|\theta) \qquad N_{0} \underbrace{(x|H_{0}, \Sigma_{0})}_{N} = P(\frac{x|Z_{0}}) = P(\frac{x|Z_{0}, \theta}{2}) \\ P(x|Z_{0}, \theta) = P(X|\theta) = P(Z_{0}) P(x|Z_{0}) + \cdots + P(Z_{k}) P(x|Z_{k}) \\ P(x|Z_{k}, \theta) P(Z_{k}) = \sum_{k}^{n} P(x|Z_{k}) P(Z_{k}) = \sum_{k}^{n} P(x, Z_{k}) = P(x) \\ P(x|Z_{k}, \theta) P(Z_{k}|\theta) = P(x|\theta) = P(x|Q) \\ P(x|Z_{k}, \theta) P(Z_{k}|\theta) = P(x, Z_{k}|\theta) = P(x|Q) \\ P(x|Z_{k}, \theta) P(Z_{k}|\theta) = P(x|Q) \\ \end{pmatrix}$$

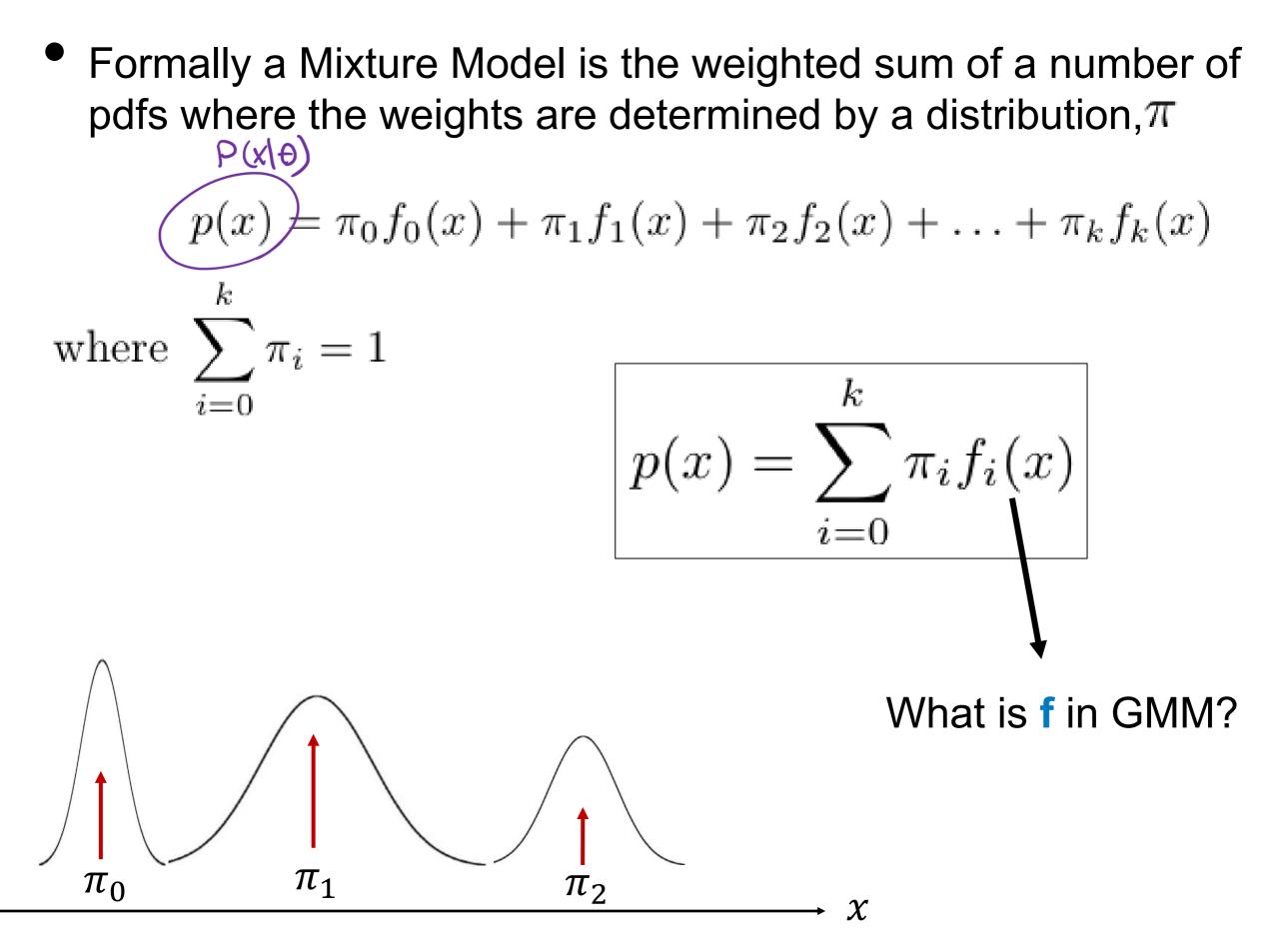
Mixture perspective – soft assignment

Mixture perspective – Initialization

 $p(\mathbf{x}) = \pi_0 N(X|\mu_0, \sigma_0) + \pi_1 N(X|\mu_1, \sigma_1) + \pi_2 N(X|\mu_2, \sigma_2)$



Mixture Models



Mixture Models are Generative

• Generative simply means dealing with joint probability p(x,z)

$$p(x) = \pi_0 f_0(x) + \pi_1 f_1(x) + \dots + \pi_k f_k(x)$$

Let's say f(.) is a Gaussian distribution

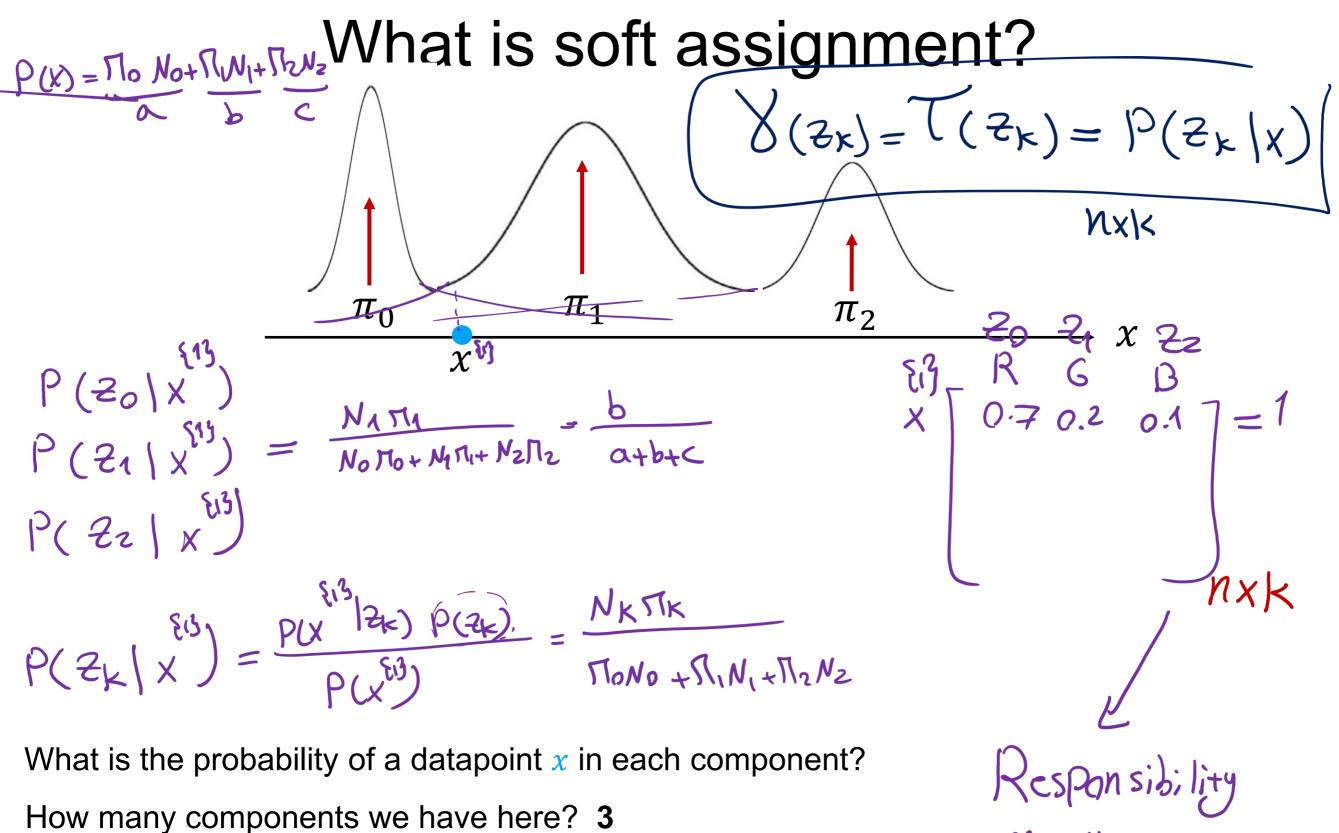
 $p(\mathbf{x}) = \pi_0 N(X|\mu_0, \sigma_0) + \pi_1 N(X|\mu_1, \sigma_1) + \dots + \pi_k N(X|\mu_k, \sigma_k)$

$$p(x) = \sum_{k} N(x|\mu_k, \sigma_k) \pi_k$$

$$p(x) = \sum_{k} p(x|z_k) p(z_k)$$

 z_k is component k

$$p(x) = \sum_{k} p(x, z_k)$$



Matrix

How many probability? 3

What is the sum value of the 3 probabilities for each datapoint? 1

Inferring Cluster Membership

- We have representations of the joint $p(x, z_{nk}|\theta)$ and the marginal, $p(x|\theta)$
- The conditional of $p(z_{nk}|x,\theta)$ can be derived using Bayes rule.
 - The responsibility that a mixture component takes for explaining an observation x.

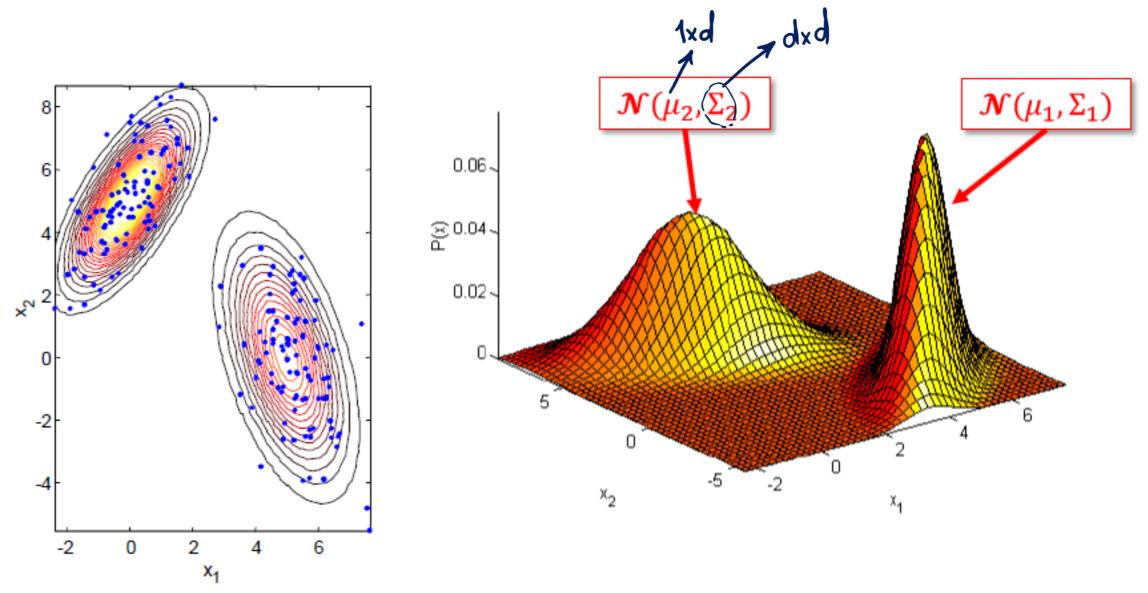
$$\tau(z_k) = p(-z_k - |x) = \frac{p(-z_k -)p(x|-z_k -)}{\sum_{j=1}^{K} p(-z_j -)p(x|-z_j -)}$$
$$= \frac{\pi_k N(x|\mu_k, \Sigma_k)}{\sum_{j=1}^{K} \pi_j N(x|\mu_j, \Sigma_j)}$$

Why having "Latent variable"

- A variable can be unobserved (latent) because:
 - it is an imaginary quantity meant to provide some simplified and abstractive view of the data generation process.
 - e.g., speech recognition models, mixture models (soft clustering)...
 - it is a real-world object and/or phenomena, but difficult or impossible to measure
 - e.g., the temperature of a star, causes of a disease, evolutionary ancestors ...
 - it is a real-world object and/or phenomena, but sometimes wasn't measured, because of faulty sensors, etc.
 - Discrete latent variables can be used to partition/cluster data into sub-groups.
 - Continuous latent variables (factors) can be used for dimensionality reduction (factor analysis, etc).

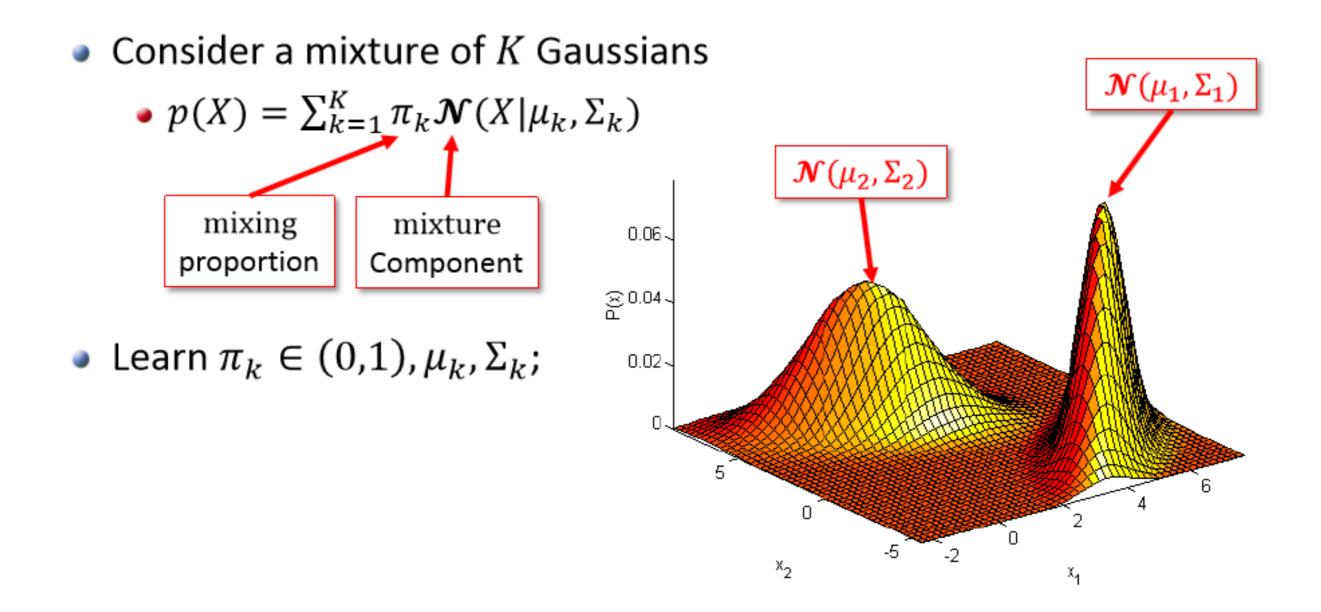
How about GMM for multimodal distribution?

- What if we know the data consists of a few Gaussians
- What if we want to fit parametric models



Gaussian Mixture Model

 A density model p(X) may be multi-modal: model it as a mixture of uni-modal distributions (e.g. Gaussians)



What are GMM parameters?

Mean μ_k Size π_k Variance σ_k

Marginal probability distribution

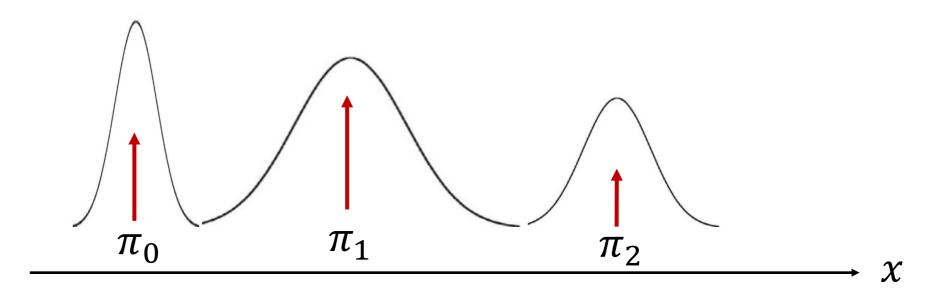
$$p(\mathbf{x}|\theta) = \sum_{k} p(x, z_{k}|\theta) = \sum_{k} p(x|z_{k}, \theta) p(z_{k}|\theta) = \sum_{k} N(x|\mu_{k}, \sigma_{k})\pi_{k}$$

$$p(z_k|\theta) = \pi_k$$
 Select a
 $p(x|z_k, \theta) = N(x|\mu_k, \sigma_k)$ Same

1

mixture component with probability π

ple from that component's Gaussian



Parameters' definition

- Purpose: GMM is a clustering algorithm derived from probabilistic theory that uses soft-assignment, meaning that data points have probability of being associated/generated from K gaussians/clusters. This is as opposed to K-means where data points definitively are either from a cluster or they're not.
- Gaussian Parameters
 - μ: Mean of each gaussian, can be compared to the K-means cluster centers
 - Σ : Covariance matrix of each gaussian, which represents how dimensions vary between each other. If it's the covariance of a specific dimension with itself, it is just the standard deviation of that variable/dimension. This is in a DxD matrix (for each gaussian and every element Σ i,j represents the covariance of dimension I with j. If you assume the dimensions are independent, then only the diagonals are non-zero.
 - z_{nk}: Latent variable which isn't explicitly known, but tells us which gaussian each datapoint was generated from.
 z is binary, it either's 1 (point x came from gaussian k) or 0 (point x did not come from gaussian k)
 - $p(z_{nk}) = \pi$: Mixing proportions/weights, which represent the fraction of data points that are generated from/associated with each gaussian. These sum to 1.
- $N(X_n | \mu_k, \Sigma_k)$: This term is the probability of some data point X_n occurring based on the assumption that it is generated from gaussian k. Mathematically, this Is equal to the likelihood ($p(x|z,\mu_k, \Sigma_k)$). Multiply this with the mixing weight π , and you get the joint distribution $p(x=X_n,z=k)$.
- P(X): If you sum up all the N(X_n | μ_k , Σ_k)^{*} π terms for each cluster, you get the probability of the entire data set occurring.
- $\gamma(z_{nk})$ or $P(z_{nk} | x_n)$: We call this term the "responsibility." It is the probability of z for "A" data point x, meaning that this is probability that point n is generated from gaussian k normalized by P(X), the probability of the entire data set occurring.

Well, we don't know π_k, μ_k, Σ_k What should we do?

We use a method called "Maximum Likelihood Estimation" (MLE) to solve the problem.

$$p(\mathbf{x}) = p(\mathbf{x}|\theta) = \sum_{k} p(x, z_k|\theta) = \sum_{k} p(z_k|\theta) p(x|z_k, \theta) = \sum_{k=0}^{K} \pi_k N(x|\mu_k, \Sigma_k)$$

Let's identify a likelihood function, why?

Because we use likelihood function to optimize the probabilistic model parameters!

$$\arg\max p(x|\theta) = p(x|\pi,\mu,\Sigma) = \prod_{n=1}^{N} p(x_n|\theta) = \prod_{n=1}^{N} \sum_{k=0}^{K} \pi_k N(x_n|\mu_k,\Sigma_k)$$

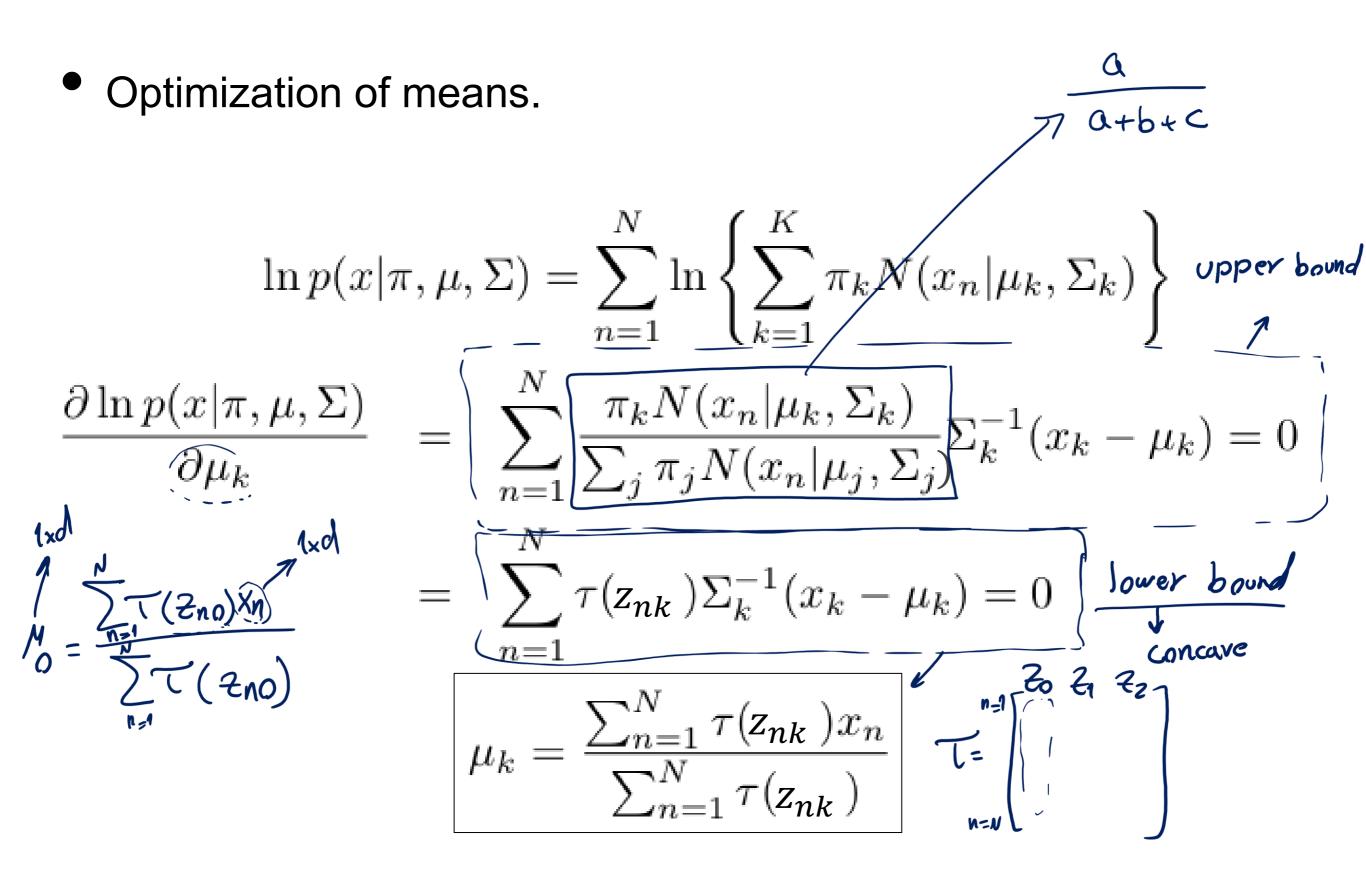
$$\arg\max p(x) = p(x|\pi, \mu, \Sigma) = \prod_{n=1}^{N} p(x_n|\theta) = \prod_{n=1}^{N} \sum_{k=0}^{K} \pi_k N(x_n|\mu_k, \Sigma_k)$$

 $\ln[p(x)] = \ln[p(x|\pi,\mu,\Sigma)]$

• As usual: Identify a likelihood function

$$\ln p(x|\pi,\mu,\Sigma) = \sum_{n=1}^{N} \ln \left\{ \sum_{k=1}^{K} \pi_k N(x_n|\mu_k,\Sigma_k) \right\}$$

Maximum Likelihood of a GMM



Maximum Likelihood of a GMM

Optimization of covariance

$$\ln p(x|\pi,\mu,\Sigma) = \sum_{n=1}^{N} \ln \left\{ \sum_{k=1}^{K} \pi_k N(x_n|\mu_k,\Sigma_k) \right\}$$

$$\Sigma_{k} = \frac{1}{\sum_{n=1}^{N} \tau(z_{nk})} \sum_{n=1}^{N} \tau(z_{nk}) (x_{n} - \mu_{k}) (x_{n} - \mu_{k})^{T}$$

Maximum Likelihood of a GMM

Optimization of mixing term

$$\ln p(x|\pi,\mu,\Sigma) + \lambda \left(\sum_{k=1}^{K} \pi_{k} - 1\right)$$

$$\Pi_{0} = \sum_{n=1}^{N} \frac{N(x_{n}|\mu_{k},\Sigma_{k})}{\sum_{j} \pi_{j}N(x_{n}|\mu_{j},\Sigma_{j})} + \lambda$$

$$\Pi_{0} = \frac{\sum_{n=1}^{N} \frac{N(x_{n}|\mu_{k},\Sigma_{k})}{\sum_{j} \pi_{j}N(x_{n}|\mu_{j},\Sigma_{j})} + \lambda$$

$$\pi_{k} = \frac{\sum_{n=1}^{N} \tau(z_{nk})}{N = 400} = \left(\sum_{n=N}^{N-1} \frac{1}{\sqrt{2}}\right) = 0$$

MLE of a GMM

$$\mu_k = \frac{\sum_{n=1}^N \tau(z_{nk}) x_n}{N_k}$$

$$\sum_{k} = \frac{1}{N_k} \sum_{n=1}^{N} \tau(z_{nk}) (x_n - \mu_k) (x_n - \mu_k)^T$$

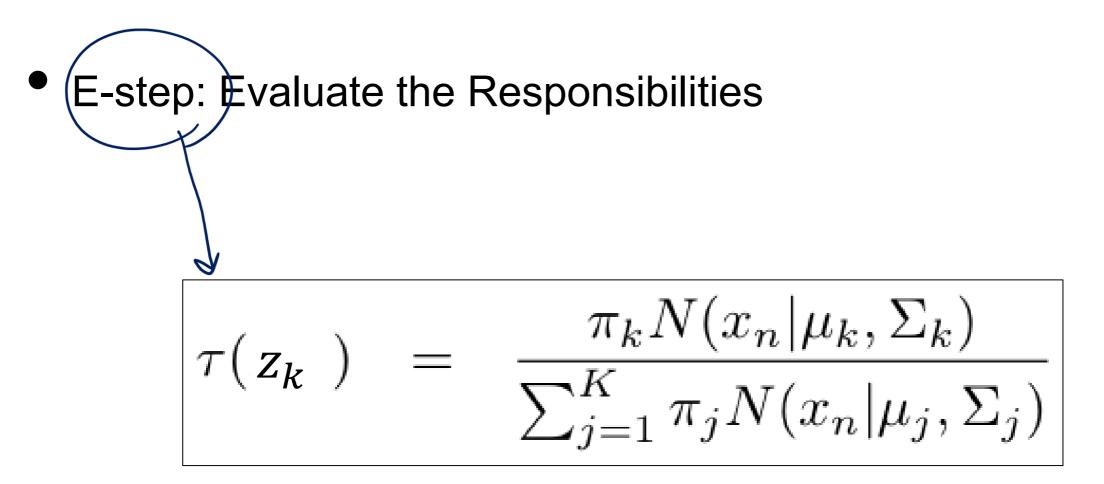
$$\pi_k = \frac{N_k}{N}$$

$$N_k = \sum_{n=1}^N \tau(z_{nk})$$

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EM for GMMs



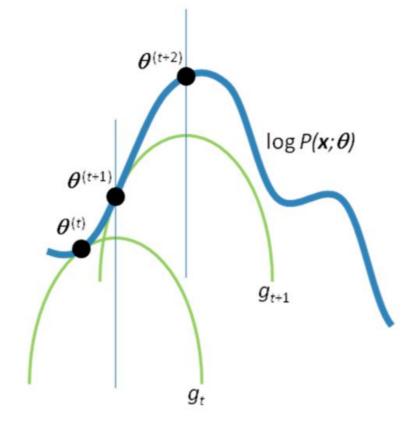
 $\mu_k^{new} = \frac{\sum_{n=1}^N \tau(z_{nk}) x_n}{N_k}$

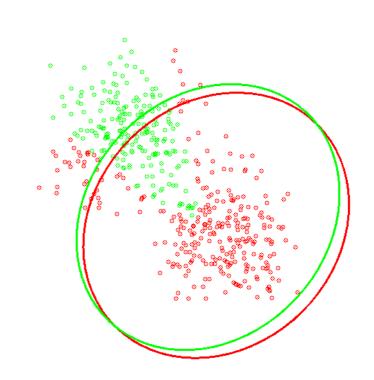
$$\sum_{k}^{new} = \frac{1}{N_k} \sum_{n=1}^{N} \tau(z_{nk}) (x_n - \mu_k^{new}) (x_n - \mu_k^{new})^T$$

$$\pi_k^{new} = \frac{N_k}{N}$$

Expectation Maximization

- Expectation Maximization (EM) is a general algorithm to deal with hidden variables.
- Two steps:
 - 。 E-Step: Fill-in hidden values using inference
 - ^o M-Step: Apply standard MLE method to estimate parameters
- EM always converges to a local minimum of the likelihood.

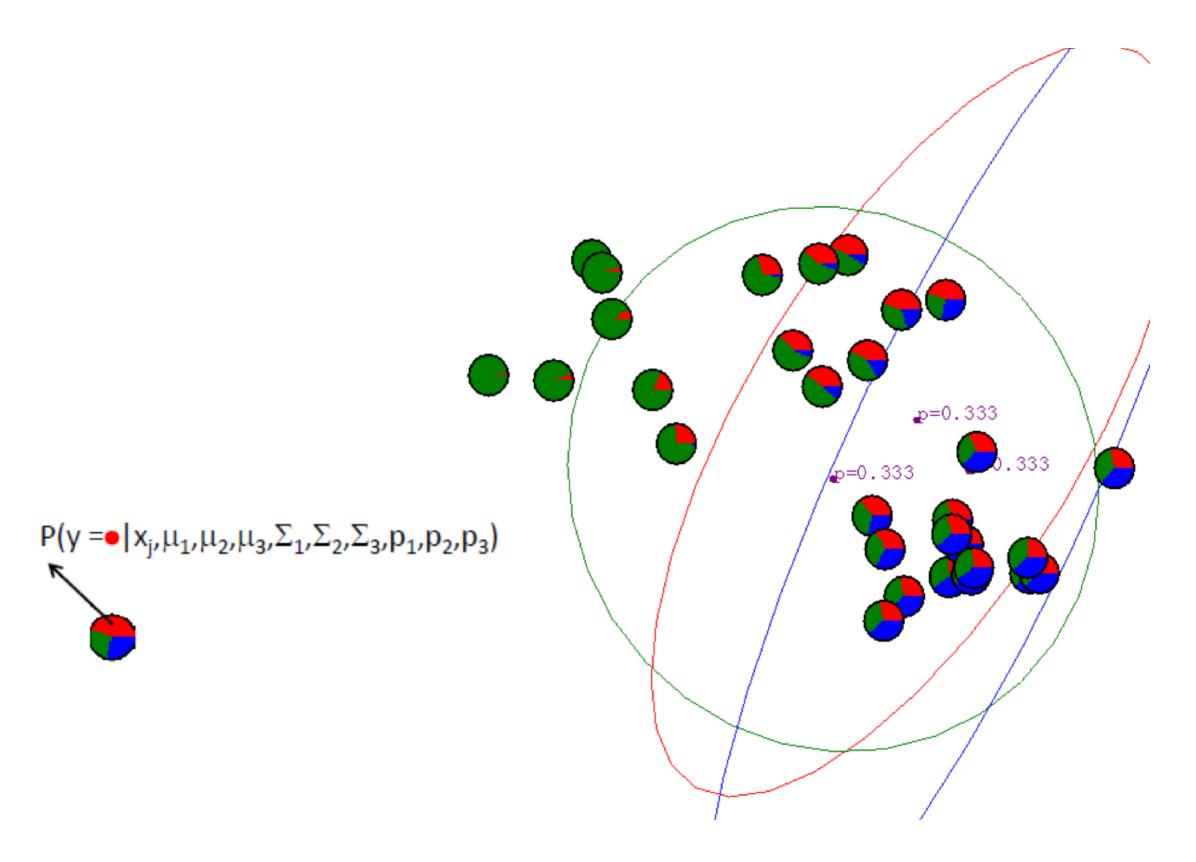




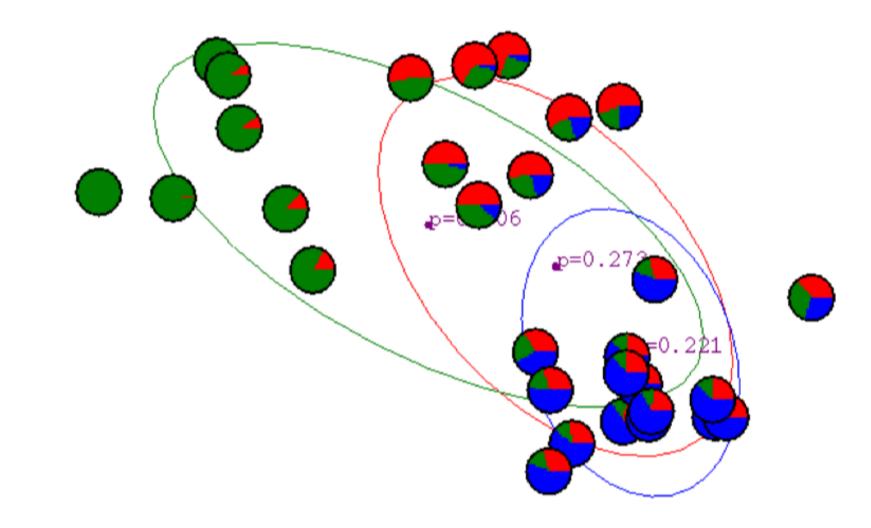
covariance_type="diag"
$$\Sigma = \begin{bmatrix} \alpha & 0 \\ 0 & b \end{bmatrix} = \begin{bmatrix} 2 & -1 \\ 0 & b \end{bmatrix} = \begin{bmatrix} -1 \\ 0 & b \end{bmatrix} = \begin{bmatrix} -1 \\ 0 & b \end{bmatrix}$$

covariance_type="spherical"

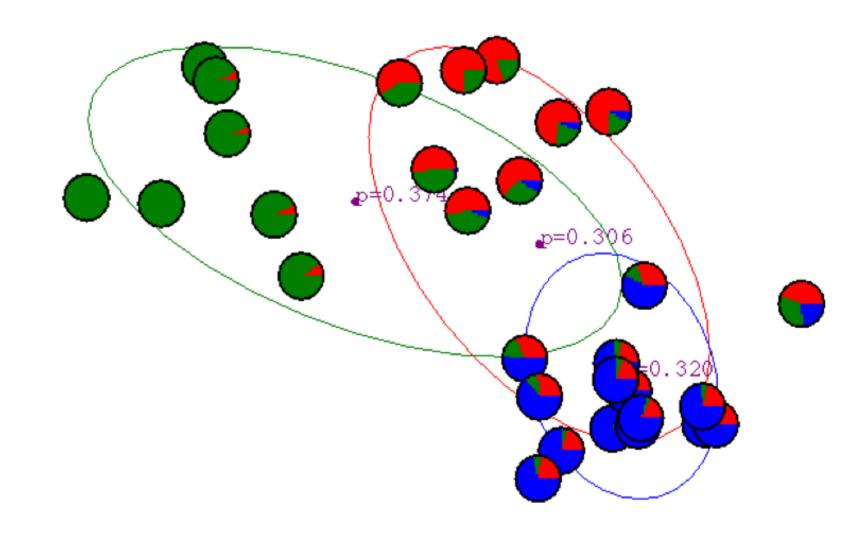
EM for Gaussian Mixture Model:



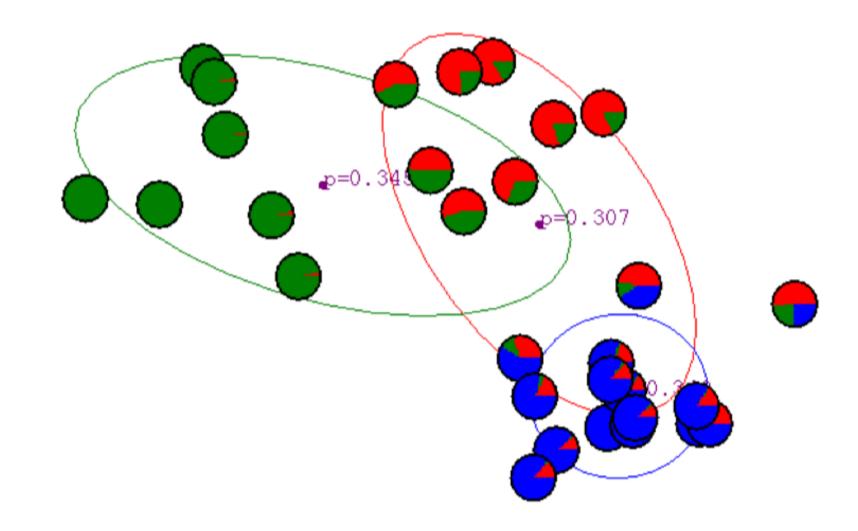
After 1st iteration



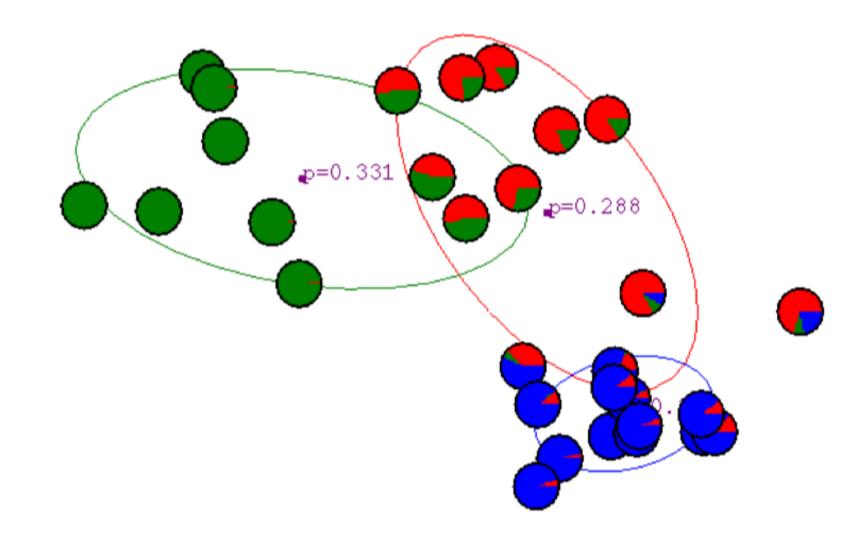
After 2nd iteration



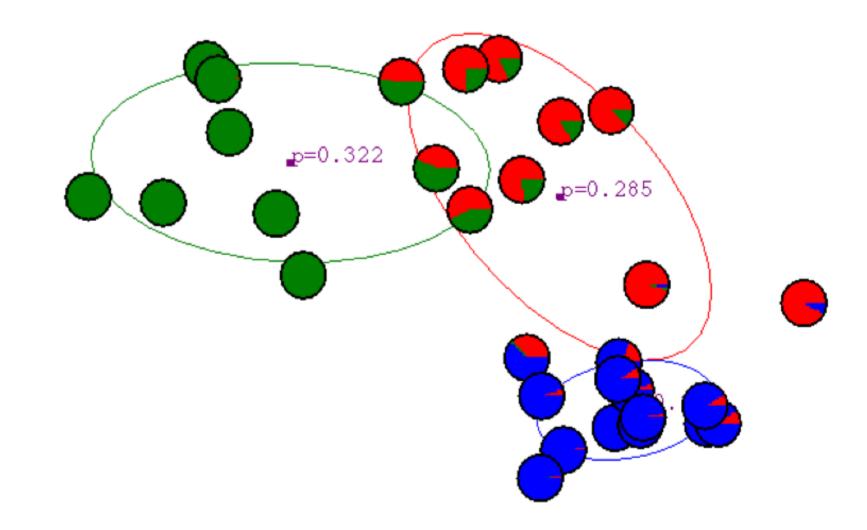
After 3rd iteration



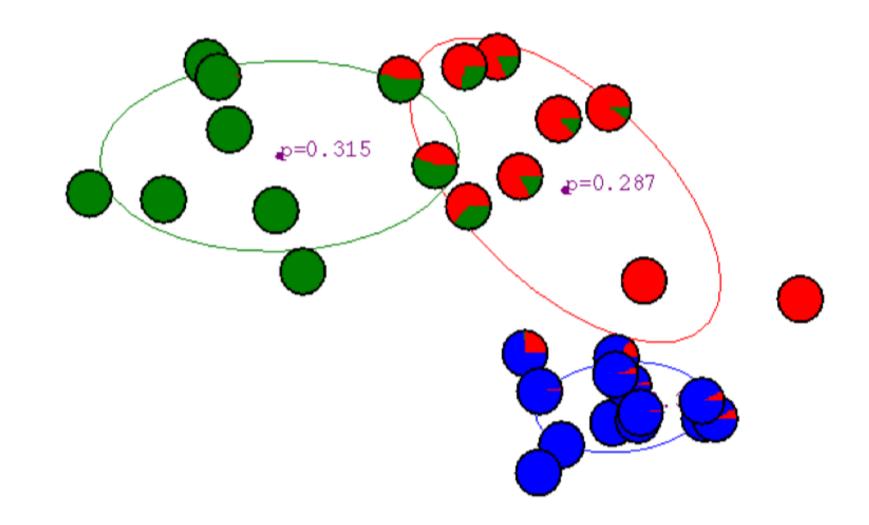
After 4th iteration



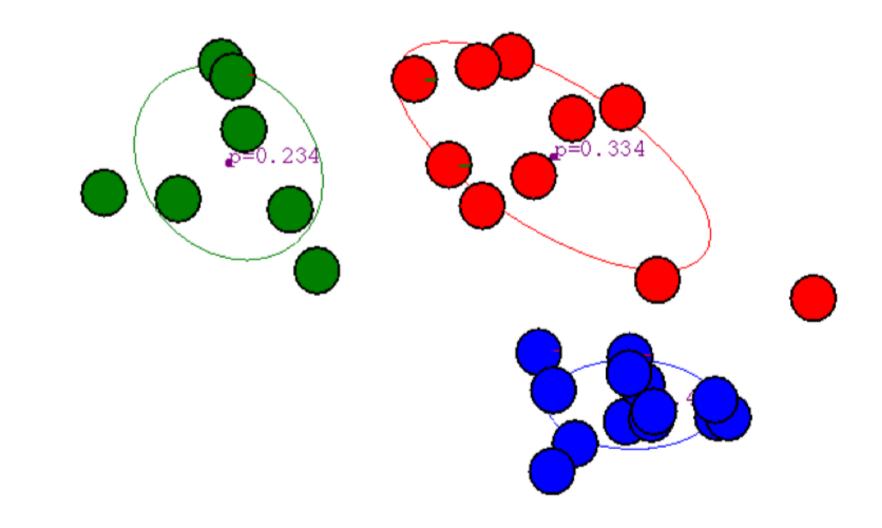
After 5th iteration



After 6th iteration



After 20th iteration



EM Algorithm for GMM (matrix form)

Given a Gaussian mixture model, the goal is to maximize the likelihood function with respect to the parameters comprising the means and covariances of the components and the mixing coefficients).

- 1. Initialize the means μ_{j} , covariances \sum_{j} and mixing coefficients π_{j} , and evaluate the initial value of the log likelihood.
- 2. E step. Evaluate the responsibilities using the current parameter values

$$\gamma(z_k) = \frac{\pi_k \mathcal{N}(\mathbf{x} \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^{K} \pi_j \mathcal{N}(\mathbf{x} \mid \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}$$

Book : C.M. Bishop, Pattern Recognition and Machine Learning, Springer, 2006

EM Algorithm for GMM (matrix form)

3. M step. Re-estimate the parameters using the current responsibilities

$$\mu_{k} = \frac{\sum_{n=1}^{N} \gamma(z_{nk}) \mathbf{X}_{n}}{\sum_{n=1}^{N} \gamma(z_{nk})} \sum_{n=1}^{N} \gamma(z_{nk}) \left(\sum_{k=1}^{N} \sum_{n=1}^{N} \gamma(z_{nk}) \sum_{n=1}^{N} \gamma(z_{nk}) \right)^{T} \left(\pi_{k} = \frac{1}{N} \sum_{n=1}^{N} \gamma(z_{nk}) \right)^{T}$$

4. Evaluate log likelihood

$$\ln p(\mathbf{X} \mid \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) = \sum_{n=1}^{N} \ln \left\{ \sum_{k=1}^{K} \pi_{k} \mathbf{N}(\mathbf{x}_{n} \mid \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}) \right\}$$

If there is no convergence, return to step 2.

Relationship to K-means

- K-means makes hard decisions.
 - 。Each data point gets assigned to a single cluster.
- GMM/EM makes soft decisions.
 - . Each data point can yield a posterior p(z|x)
- K-means is a special case of EM.

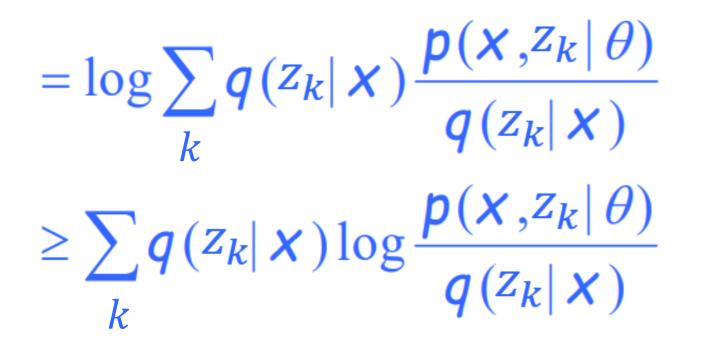
General form of EM

- Given a joint distribution over observed and latent variables: $p(X, Z|\theta)$
- Want to maximize: $p(X|\theta)$
- 1. Initialize parameters θ^{old}
- 2. E Step: Evaluate: $p(Z|X, \theta^{old})$
- 3. M-Step: Re-estimate parameters (based on expectation of completedata log likelihood)

$$\theta^{new} = \operatorname{argmax}_{\theta} \sum_{Z} p(Z_k | X, \theta^{old}) \ln p(X, Z_k | \theta) = \operatorname{argmax}_{\theta} Exp[\log(p(x, Z_k | \theta))]$$

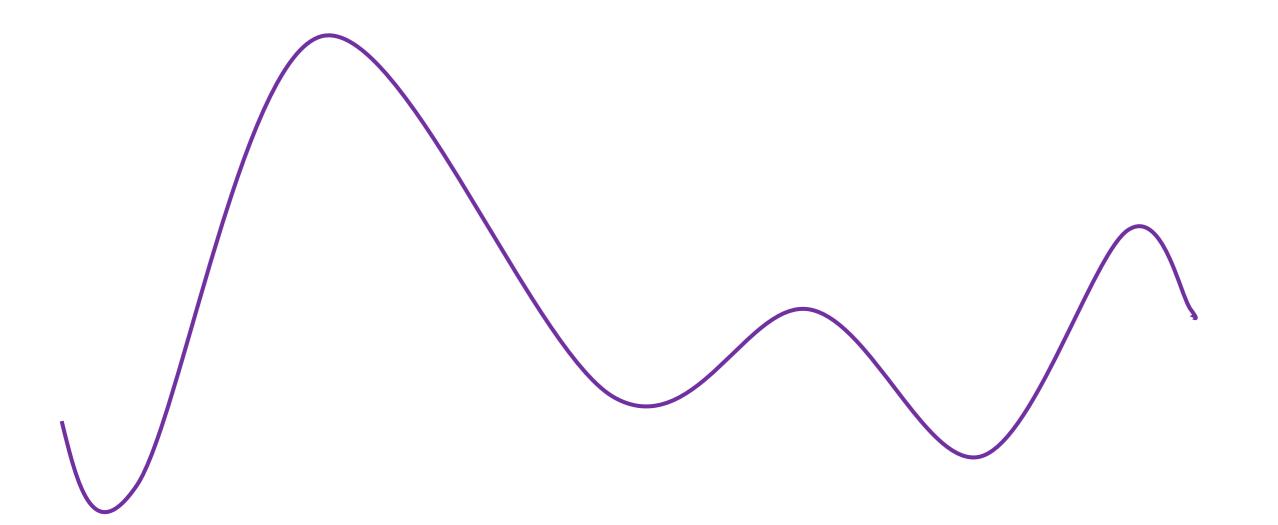
1. Check for convergence of params or likelihood

$$\theta^{new} = \operatorname{argmax}_{\theta} \sum_{k} p(\bar{z_k} | X, \theta^{old}) \ln p(X, \bar{z_k} | \theta)$$

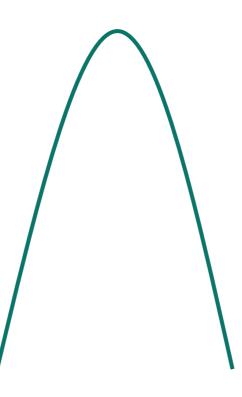


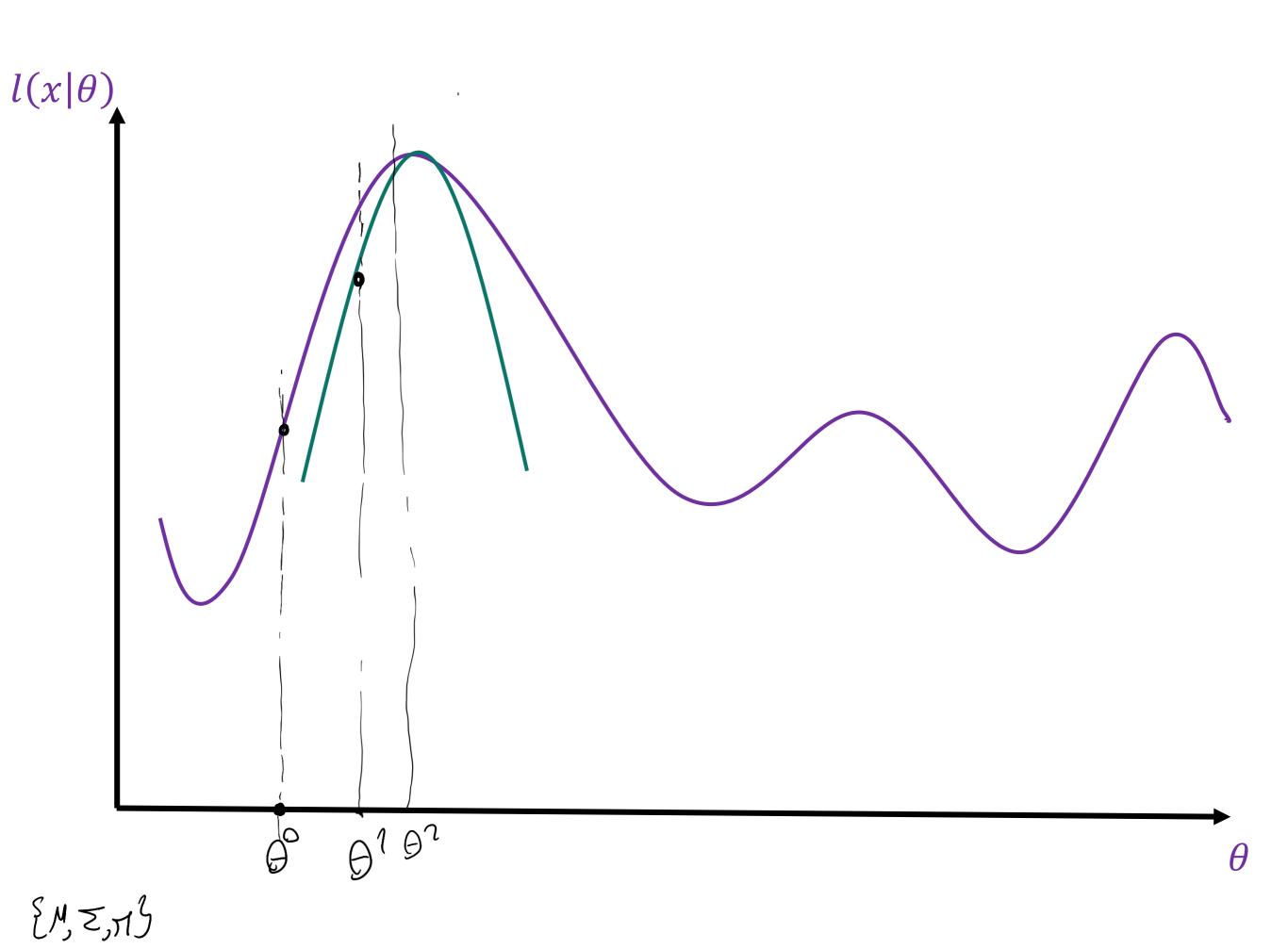
$$l(\theta|x) = \log p(x|\theta) = \log \sum_{k} p(x, z_{k}|\theta) \ge \sum_{k} q(z_{k}|x) \log \frac{p(x, z_{k}|\theta)}{q(z_{k}|x)}$$
$$\log \left(\sum_{k} p(x, z_{k}|\theta)\right) = \log \left(\sum_{k} p(x|\theta, z_{k}) * p(z_{k}|\theta)\right)$$

$$= \log(N(x|\mu_0, \Sigma_0) * \pi_0 + \dots + N(x|\mu_k, \Sigma_k) * \pi_0)$$



$$\begin{aligned} \mathcal{L}(\theta|x) &= \log p(x|\theta) = \log \sum_{k} p(x, z_{k}|\theta) \geq \sum_{k} q(z_{k}|x) \log \frac{p(x, z_{k}|\theta)}{q(z_{k}|x)} \\ q(z_{k}|x) &= C_{k} \Rightarrow \text{It is given to us} \\ \sum_{k} q(z_{k}|x) \log \frac{p(x, z_{k}|\theta)}{q(z_{k}|x)} = \\ C_{0} \log \left(\frac{1}{C_{0}} * N(x|\mu_{0}, \Sigma_{0}) * \pi_{0}\right) + \dots + C_{k} \log \left(\frac{1}{C_{k}} * N(x|\mu_{k}, \Sigma_{k}) * \pi_{k}\right) \end{aligned}$$





$$l(\theta|x) = \log p(x|\theta) = \log \sum_{z} p(x,z|\theta) \stackrel{\text{obs}}{=} \sum_{z} q(z|x) \log \frac{p(x,z|\theta)}{q(z|x)} \stackrel{\text{hs}}{=} \log p(x|\theta) = \log p(x|\theta) \stackrel{\text{hs}}{=} \log p(x|\theta) = \log p(x|x|\theta) + \log p(x|x|\theta)$$

$$\frac{q(z|x)}{q(z|x)} \stackrel{\text{hs}}{=} \log p(z|x,\theta) \stackrel{\text{hs}}{=} \log p(x|\theta) - \log p(z|x,\theta) \stackrel{\text{hs}}{=} \log p(z|x,\theta) \stackrel{\text{hs}}{=} \log p(z|x,\theta) \stackrel{\text{hs}}{=} \log p(x|\theta) - \log p(z|x,\theta) \stackrel{\text{hs}}{=} \log p(x|\theta) \stackrel{$$

$$\ell(heta|x) = \sum_{k\in\{0,1\}}^N \sum_{k\in\{0,1\}}^Z p(z_k|x_n, heta_{old}) ln\left[p(x_n,z_k| heta)
ight]$$

$$\log \sum_{k} P(x, 2k|\theta) \ge \sum_{k} q(2k|x) \log \frac{P(x, 2k|\theta)}{P(2k|x)}$$

$$If \Rightarrow q(2k|x) = P(2k|x, 0^{old})$$

$$lower bound = \sum_{k} P(2k|x, 0^{old}) \log \frac{P(x, 2k|\theta)}{P(2k|x, 0^{old})}$$

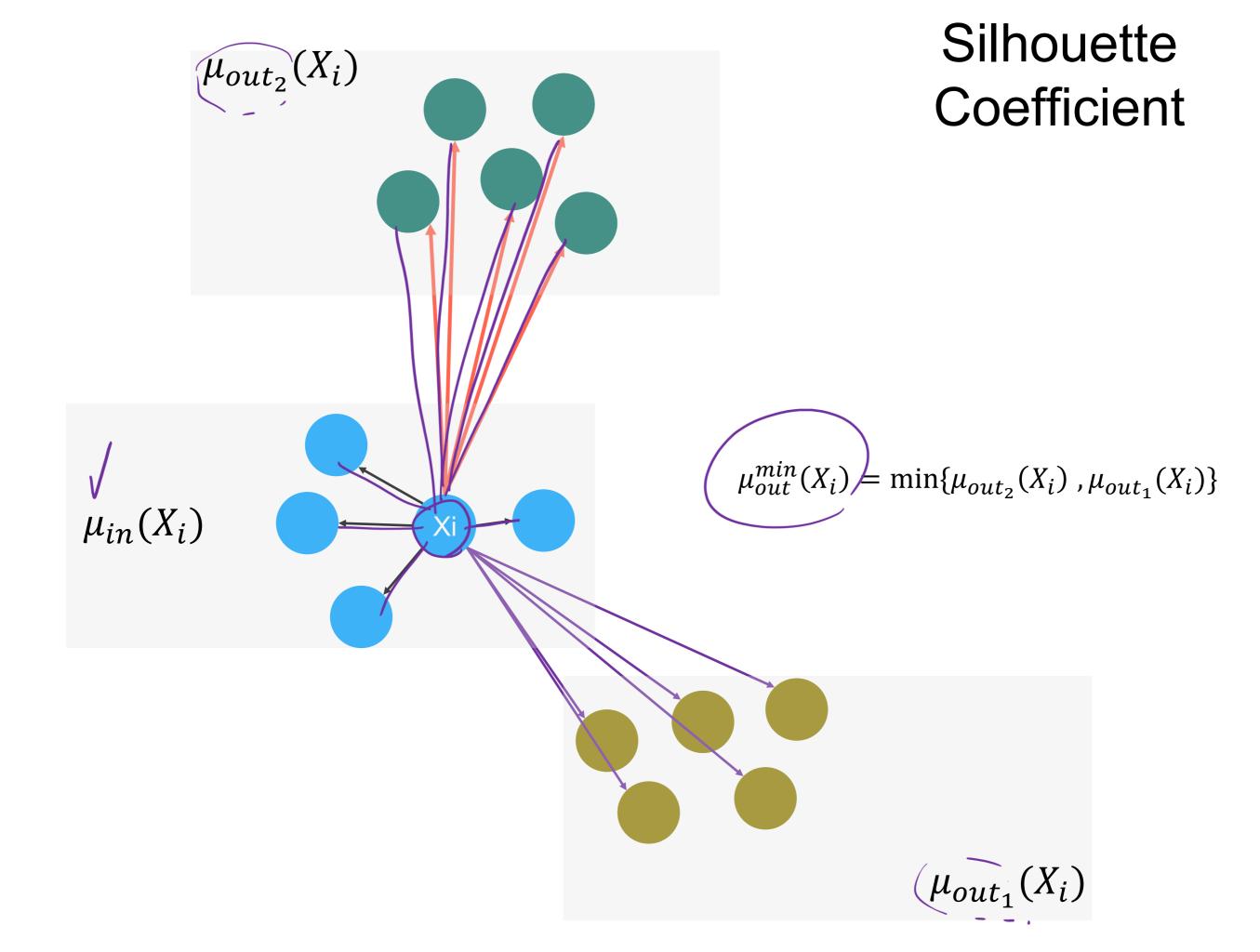
$$lower bound = \sum_{k} P(2k|x, 0^{old}) \log \frac{P(x, 2k|\theta)}{P(2k|x, 0^{old})}$$

$$lower bound = \sum_{k} P(2k|x, 0^{old}) \log P(x, 2k|\theta) - \log P(2k|x, 0^{old})$$

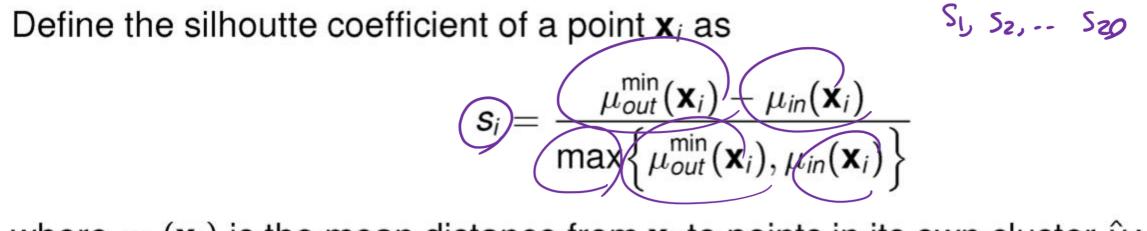
$$lower bound = \sum_{k} P(2k|x, 0^{old}) \log P(x, 2k|\theta) + \sum_{k} P(2k|x, 0^{old}) \times \frac{1}{163} P(2k|x, 0^{old}) \times \frac{1}{163} P(2k|x, 0^{old})$$

$$lower bound = \sum_{k} P(2k|x, 0^{old}) \log P(x, 2k|\theta) + \sum_{k} P(2k|x, 0^{old}) \times \frac{1}{163} P(2k|x, 0^{old}) \times \frac{1}{163}$$

<u>https://www.dropbox.com/</u> <u>scl/fi/j0nmxf654bbluf3zowp</u> <u>78/EM-maximization-and-</u> <u>equality.mp4?rlkey=wlua7l</u> <u>5r88kdtydjoru7qc6f6&st=5</u> <u>h7rbts6&dI=0</u>



Silhouette Coefficient



where $\mu_{in}(\mathbf{x}_i)$ is the mean distance from \mathbf{x}_i to points in its own cluster \hat{y}_i :

$$\mu_{in}(\mathbf{x}_i) = \frac{\sum_{\mathbf{x}_j \in C_{\hat{y}_i}, j \neq i} \delta(\mathbf{x}_i, \mathbf{x}_j)}{n_{\hat{y}_i} - 1}$$

and $\mu_{out}^{\min}(\mathbf{x}_i)$ is the mean of the distances from \mathbf{x}_i to points in the closest cluster:

$$\mu_{out}^{\min}(\mathbf{x}_i) = \min_{j \neq \hat{y}_i} \left\{ \frac{\sum_{\mathbf{y} \in C_j} \delta(\mathbf{x}_i, \mathbf{y})}{n_j} \right\}$$

The Silhouette Coefficient for clustering C: SC =

$$C=\frac{1}{n}\sum_{i=1}^{n}\mathbf{s}_{i}.$$

SC close to 1 implies a good clustering (Points are close to their own clusters but far from other clusters)

Take-Home Messages

- The generative process of Gaussian Mixture Model
- Inferring cluster membership based on a learned GMM
- The general idea of Expectation-Maximization
- Expectation-Maximization for GMM
- Silhouette Coefficient